# Flerida <br> Standards Assessments 

# Florida Standards Assessments 

## 2017-2018

# Volume 4 <br> Evidence of Reliability and Validity 

FLORIDA DEPARTMENT OF


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## 1. Introduction and Overview of Reliability and Validity Evidence

The State of Florida implemented a new assessment program for operational use during the 20142015 school year. This new program, named the Florida Standards Assessments (FSA), replaced the Florida Comprehensive Assessment Tests (FCAT) 2.0 in English Language Arts (ELA) and Mathematics. Students in grades 3 and 4 were administered fixed, operational ELA Reading and Mathematics tests on paper. Students in grades 5 through 10 were administered fixed, operational Reading tests online, and students in grades 5 through 8 were administered fixed, operational Mathematics tests online. Online End-of-Course (EOC) assessments were administered to students taking Algebra 1, Algebra 2, and Geometry. In addition, students in grades 4 through 10 responded to a text-based Writing prompt, with grades 4 through 7 administered on paper, and grades 8 through 10 administered online. Writing and Reading scores were combined to form an overall ELA score. In the spring of 2016, the grade 4 Reading portion of the ELA assessment transition to an online delivery. In the spring of 2017, the grades 3 and 4 Mathematics assessment moved online, too. Beginning in summer 2017, Algebra 2 was no longer administered.

In the grades with online testing, paper forms, in lieu of online forms, were administered to students whose Individual Educational Plans (IEP) or Section 504 Plans indicated such a need. Grade 3 Reading was universally administered on paper, so there was no accommodated form. Table 1 displays the complete list of test forms for the 2017-2018 spring operational administration.

Table 1: Test Administration

| Subject | Administration | Grade/Course |
| :--- | :--- | :--- |
| ELA Reading | Paper | 3 |
| ELA Reading | Online | $4-10$ |
|  | Paper (Accommodated) | $4-7$ |
| ELA Writing | Paper | $8-10$ |
|  | Online |  |
|  | Paper (Accommodated) | $3-8$ |
| EOC | Online | Algebra 1, Geometry |
|  | Paper (Accommodated) | Online |

With the implementation of these tests, both reliability evidence and validity evidence are necessary to support appropriate inferences of student academic achievement from the FSA scores. This volume provides empirical evidence about the reliability and validity of the 2017-2018 FSA, given its intended uses.

The purpose of this volume is to provide empirical evidence to support the following:

- Reliability: Multiple reliability estimates for each test are reported in this volume, including stratified-coefficient alpha, Feldt-Raju, and the marginal reliability. The reliability estimates are presented by grade and subject as well as by demographic subgroups. This section also includes conditional standard errors of measurement and classification accuracy results by grade and subject.
- Content validity: Evidence is provided to show that test forms were constructed to measure the Florida Standards with a sufficient number of items targeting each area of the blueprint.
- Internal structure validity: Evidence is provided regarding the internal relationships among the subscale scores to support their use and to justify the item response theory (IRT) measurement model. This type of evidence includes observed and disattenuated Pearson correlations among reporting categories per grade. Confirmatory factor analysis has also been performed using the second-order factor model. Additionally, local item independence, an assumption of unidimensional IRT, was tested using the $\mathrm{Q}_{3}$ statistic.
- Comparability of paper-pencil to online tests: By examining the blueprint match between forms and test characteristic curves (TCCs) for both forms, we evaluate comparability of test scores across forms.
- Test fairness: Fairness is statistically analyzed using differential item functioning (DIF) in tandem with content alignment reviews by specialists.


### 1.1 RELIABILITY

Reliability refers to consistency in test scores. Reliability can be defined as the degree to which individuals' deviation scores remain relatively consistent over repeated administrations of the same test or alternate test forms (Crocker \& Algina, 1986). For example, if a person takes the same or parallel tests repeatedly, he or she should receive consistent results. The reliability coefficient refers to the ratio of true score variance to observed score variance:

$$
\rho_{\mathrm{XX}}{ }^{\prime}=\frac{\sigma_{\mathrm{T}}^{2}}{\sigma_{\mathrm{X}}^{2}}
$$

There are various approaches for estimating the reliability of scores. The conventional approaches used are characterized as follows:

- The test-retest method measures stability over time. With this method, the same test is administered twice to the same group at two different points in time. If test scores from the two administrations are highly correlated, then the test scores are deemed to have a high level of stability. For example, if the result is highly stable, those who scored high on the first administration tend to obtain a high score on the second administration. The critical factor, however, is the time interval. The time interval should not be too long, which could allow for changes in the examinees' true scores. Likewise, it should not be too short, in which case memory and practice may confound the results. The test-retest method is most effective for measuring constructs that are stable over time, such as intelligence or personality traits.
- The parallel-forms method is used for measuring equivalence. With this design, two parallel forms of the test are administered to the same group. This method requires two
similar forms of a test. However, it is very difficult to create two strictly parallel forms. When this method is applied, the effects of memory or practice can be eliminated or reduced, since the tests are not purely identical as with the test-retest method. The reliability coefficient from this method indicates the degree to which the two tests are measuring the same construct. While there are a wide variety of possible items to administer to measure any particular construct, it is only feasible to administer a sample of items on any given test. If there is a high correlation between the scores of the two tests, then inferences regarding high reliability of scores can be substantiated. This method is commonly used to estimate the reliability of achievement or aptitude tests.
- The split-half method utilizes one test divided into two halves within a single test administration. It is crucial to make the two half-tests as parallel as possible, as the correlation between the two half-tests is used to estimate reliability of the whole test. In general, this method produces a coefficient that underestimates the reliability for the full test. To correct the estimate, the Spearman-Brown prophecy formula (Brown, 1910; Spearman, 1910) can be applied. While this method is convenient, varying splits of the items may yield different reliability estimates.
- The internal consistency method can be employed when it is not possible to conduct repeated test administrations. Whereas other methods often compute the correlation between two separate tests, this method considers each item within a test to be a one-item test. There are several other statistical methods based on this idea: coefficient alpha (Cronbach, 1951), Kuder-Richardson Formula 20 (Kuder \& Richardson, 1937), KuderRichardson Formula 21 (Kuder \& Richardson, 1937), stratified coefficient alpha (Qualls, 1995), and Feldt-Raju coefficient (Feldt \& Qualls, 1996; Feldt \& Brennan, 1989).
- Inter-rater reliability is the extent to which two or more individuals (coders or raters) agree. Inter-rater reliability addresses the consistency of the implementation of a rating system.

Another way to view reliability is to consider its relationship with the standard errors of measurement (SEM) - the smaller the standard error, the higher the precision of the test scores. For example, classical test theory assumes that an observed score $(X)$ of each individual can be expressed as a true score $(T)$ plus some error $(E), X=T+E$. The variance of $X$ can be shown to be the sum of two orthogonal variance components:

$$
\sigma_{X}^{2}=\sigma_{T}^{2}+\sigma_{E}^{2}
$$

Returning to the definition of reliability as the ratio of true score variance to observed score variance, we can arrive at:

$$
\rho_{\mathrm{XX}}{ }^{\prime}=\frac{\sigma_{\mathrm{T}}^{2}}{\sigma_{\mathrm{X}}^{2}}=\frac{\sigma_{\mathrm{x}}^{2}-\sigma_{\mathrm{E}}^{2}}{\sigma_{\mathrm{X}}^{2}}=1-\frac{\sigma_{\mathrm{E}}^{2}}{\sigma_{\mathrm{X}}^{2}}
$$

As the fraction of error variance to observed score variance tends to zero, the reliability then tends to 1 . The Classical Test Theory (CTT) SEM, which assumes a homoscedastic error, is derived
from the classical notion expressed earlier as $\sigma_{X} \sqrt{1-\rho_{\mathrm{XX}}}$,, where $\sigma_{X}$ is the standard deviation of the scaled score and $\rho_{\mathrm{XX}}$ is a reliability coefficient. Based on the definition of reliability, this formula can be derived.

$$
\begin{gathered}
\rho_{\mathrm{XX}}=1-\frac{\sigma_{\mathrm{E}}^{2}}{\sigma_{\mathrm{X}}^{2}} \\
\frac{\sigma_{\mathrm{E}}^{2}}{\sigma_{\mathrm{X}}^{2}}=1-\rho_{\mathrm{XX}^{\prime}}, \\
\sigma_{\mathrm{E}}^{2}=\sigma_{\mathrm{X}}^{2}\left(1-\rho_{\mathrm{XX}^{\prime}}\right), \\
\sigma_{E}=\sigma_{X} \sqrt{\left(1-\rho_{\mathrm{XX}^{\prime}}\right)}
\end{gathered}
$$

In general, the standard error of measurement is relatively constant across samples as the group dependent term, $\sigma_{X}$, can be shown to cancel out:

$$
\left.\sigma_{E}=\sigma_{X} \sqrt{\left(1-\rho_{\mathrm{XX}}{ }^{\prime}\right.}\right)=\sigma_{X} \sqrt{\left(1-\left(1-\frac{\sigma_{\mathrm{E}}^{2}}{\sigma_{\mathrm{X}}^{2}}\right)\right)}=\sigma_{X} \sqrt{\frac{\sigma_{\mathrm{E}}^{2}}{\sigma_{\mathrm{X}}^{2}}}=\sigma_{X} \cdot \frac{\sigma_{E}}{\sigma_{X}}=\sigma_{E}
$$

This shows that the standard error of measurement in the classical test theory is assumed to be homoscedastic error irrespective of the standard deviation of a group.

In contrast, the standard errors of measurement in IRT vary over the ability continuum. These heterogeneous errors are a function of a test information function that provides different information about examinees depending on their estimated abilities. Often, the test information function (TIF) is maximized over an important performance cut, such as the proficient cut score.

Because the TIF indicates the amount of information provided by the test at different points along the ability scale, its inverse indicates the "lack" of information at different points along the ability scale. This lack of information is the uncertainty, or the measurement error, of the score at various score points. Conventionally, fixed-form tests are maximized near the middle of the score distribution, or near an important classification cut, and have less information at the tails of the score distribution. See Section 3.3 for the derivation of heterogeneous errors in IRT.

### 1.2 Validity

Validity refers to the degree to which "evidence and theory support the interpretations of test scores entailed by proposed uses of tests" (American Educational Research Association [AERA], American Psychological Association [APA], and National Council on Measurement in Education [NCME], 2014). Messick (1989) defines validity as "an integrated evaluative judgment of the degree to which empirical evidence and theoretical rationales support the adequacy and appropriateness of inferences and actions based on test scores and other modes of assessment." Both of these definitions emphasize evidence and theory to support inferences and interpretations of test scores. The Standards (AERA, APA, and NCME, 2014) suggests five sources of validity evidence that can be used in evaluating a proposed interpretation of test scores. When validating test scores, these sources of evidence should be carefully considered.

The first source of evidence for validity is the relationship between the test content and the intended test construct (see Section 4.4). In order for test score inferences to support a validity claim, the items should be representative of the content domain, and the content domain should be relevant to the proposed interpretation of test scores. To determine content representativeness, diverse panels of content experts conduct alignment studies, in which experts review individual items and rate them based on how well they match the test specifications or cognitive skills required for a particular construct (see Volume 2 for details). Test scores can be used to support an intended validity claim when they contain minimal construct irrelevant variance. For example, a mathematics item targeting a specific mathematics skill that requires advanced reading proficiency and vocabulary has a high level of construct irrelevant variance. Thus, the intended construct of measurement is confounded, which impedes the validity of the test scores. Statistical analyses, such as factor analysis or multi-dimensional scaling of relevance, are also used to evaluate content relevance. Results from factor analysis for the FSA are presented in Section 5.2. Evidence based on test content is a crucial component of validity, because construct underrepresentation or irrelevancy could result in unfair advantages or disadvantages to one or more group of examinees.

Technology-enhanced items should be examined to ensure that no construct irrelevant variance is introduced. If some aspect of the technology impedes, or advantages, a student in his or her responses to items, this could affect item responses and inferences regarding abilities on the measured construct. Florida makes use of the technology-enhanced items developed by AIR, and the items are delivered by the same engine as is used for delivery of the Smarter Balanced assessment. Hence, the FSA makes use of items that have the same technology-enhanced functionality as those found on these other assessments. A cognitive lab study was completed for the Smarter Balanced assessment, providing evidence in support of the item types used for the Consortium and also in Florida (see Volume 7 of the 2014-2015 Florida Standards Assessments Technical Reports).

The second source of validity evidence is based on "the fit between the construct and the detailed nature of performance or response actually engaged in by examinees" (AERA, APA, and NCME, 2014). This evidence is collected by surveying examinees about their performance strategies or responses to particular items. Because items are developed to measure particular constructs and intellectual processes, evidence that examinees have engaged in relevant performance strategies to correctly answer the items supports the validity of the test scores.

The third source of evidence for validity is based on internal structure: the degree to which the relationships among test items and test components relate to the construct on which the proposed test scores are interpreted. Differential item functioning, which determines whether particular items may function differently for subgroups of examinees, is one method for analyzing the internal structure of tests (see Volume 1, Section 5.2). Other possible analyses to examine internal structure are dimensionality assessment, goodness-of-model-fit to data, and reliability analysis (see Sections 3 and 5 for details).

A fourth source of evidence for validity is the relationship of test scores to external variables. The Standards (AERA, APA, and NCME, 2014) divides this source of evidence into three parts: convergent and discriminant evidence, test-criterion relationships, and validity generalization. Convergent evidence supports the relationship between the test and other measures intended to assess similar constructs. Conversely, discriminant evidence delineates the test from other measures intended to assess different constructs. To analyze both convergent and discriminant evidence, a multi-trait-multimethod matrix can be used. Additionally, test-criterion relationships
indicate how accurately test scores predict criterion performance. The degree of accuracy mainly depends upon the purpose of the test, such as classification, diagnosis, or selection. Test-criterion evidence is also used to investigate predictions of favoring different groups. Due to construct underrepresentation or construct-irrelevant components, the relation of test scores to a relevant criterion may differ from one group to another. Furthermore, validity generalization is related to whether the evidence is situation-specific or can be generalized across different settings and times. For example, sampling errors or range restriction may need to be considered to determine whether the conclusions of a test can be assumed for the larger population.

A study linking state tests to the National Assessment of Educational Progress (NAEP) test (Phillips, 2016) found that the Florida grades 4 and 8 level 4 performance standards, in both Mathematics and ELA, mapped to the NAEP proficiency levels. This is a rigorous standard that only Florida met as reported by Phillips (2016).

Fifth, the intended and unintended consequences of test use should be included in the testvalidation process. Determining the validity of the test should depend upon evidence directly related to the test; this process should not be influenced by external factors. For example, if an employer administers a test to determine hiring rates for different groups of people, an unequal distribution of skills related to the measurement construct does not necessarily imply a lack of validity for the test. However, if the unequal distribution of scores is in fact due to an unintended, confounding aspect of the test, this would interfere with the test's validity. As described in Volume 1 and additionally in this volume, test use should align with the intended purpose of the test.

Supporting a validity argument requires multiple sources of validity evidence. This then allows for one to evaluate if sufficient evidence has been presented to support the intended uses and interpretations of the test scores. Thus, determining the validity of a test first requires an explicit statement regarding the intended uses of the test scores, and subsequently, evidence that the scores can be used to support these inferences.

## 2. Purpose of Florida Standards Assessment

The Florida Standards Assessments (FSA) are standards-based, summative tests that measure students' achievement of Florida's education standards. Assessment supports instruction and student learning, and the results help Florida's educational leadership and stakeholders determine whether the goals of the education system are being met. Assessments help Florida determine whether it has equipped its students with the knowledge and skills they need to be ready for careers and college-level coursework. The tests are constructed to meet rigorous technical criteria outlined in Standards for Educational and Psychological Testing (AERA, APA, and NCME, 2014) and to ensure that all students have access to the test content via principles of universal design and appropriate accommodations.

The FSA yields test scores that are useful for understanding to what degree individual students have mastered the Florida Standards and, eventually, whether students are improving in their performance over time. Additionally, scores can be aggregated to evaluate the performance of subgroups, and both individual and aggregated scores will be compared over time in program evaluation methods.

The FSA results serve as the primary indicator for the state's accountability system, and the policy and legislative purpose of the FSA is described more thoroughly in Volume 1. The test is a standards-based assessment designed to measure student achievement toward the state content standards. FSA scores are indications of what students know and are able to do relative to the expectations by grade and subject area. While there are student-level stakes associated with the assessment, particularly for grade 3 ELA (scores inform district promotion decisions) and grade 10 ELA and Algebra 1 (assessment graduation requirements), the assessment is never the sole determinant in making these decisions.

Test items were selected prior to the test administration to ensure that the test construction aligned to the approved blueprint. The content and psychometric verification log was kept to track the compliance of the test structure to the FSA requirements.

In the FSA administered in 2018, student-level scores included scale scores and raw scores at the reporting category level. Based on the performance cuts approved by the State Board of Education on January 6, 2016, scale scores and achievement levels were reported in spring 2018. Volume 1 Section 8.1 of the FSA Annual Technical Report describes how each of these scores is computed.

The raw scores for reporting categories were provided for each student to indicate student strengths and weaknesses in different content areas of the test relative to the other areas and to the district and state. These scores serve as useful feedback for teachers to tailor their instruction, provided that they are viewed with the usual caution that accompanies use of reporting category scores. Thus, we must examine the reliability coefficients for these test scores and the validity of the test scores to support practical use across the state.

## 3. Reliability

### 3.1 Internal Consistency

As the FSA was administered in a single administration, it is necessary to examine the internal consistency of the test to support the reliability of the test scores. For the FSA ELA, Mathematics, and EOC assessments, the reliability coefficients were computed using Cronbach alpha, stratified alpha, and Feldt-Raju coefficient. In addition to Cronbach alpha, stratified alpha and Feldt-Raju coefficients were computed treating multiple-choice and non-multiple-choice items as two separate strata.

The FSA ELA, Mathematics, and EOC assessments included mixed item types: multiple-choice, short-response, and extended-response. Although there are various techniques for estimating the reliability of test scores with multiple item types or parts (Feldt \& Brennan, 1989; Lee \& Frisbie, 1999; Qualls, 1995), studies (Qualls, 1995; Yoon \& Young, 2000) indicate that the use of Cronbach alpha underestimates the reliability of test scores for a test with mixed item types.

The Cronbach alpha is defined as

$$
\alpha=\frac{n}{n-1}\left[1-\frac{\sum_{i=1}^{n} \sigma_{i}^{2}}{\sigma_{x}^{2}}\right],
$$

where $\sigma_{i}^{2}$ is the variance of scores on each item, $\sigma_{x}^{2}$ is the variance of the total test scores, and $n$ is the number of items.

The stratified Cronbach alpha coefficient is computed as

$$
\text { stratified } \alpha \rho_{X X^{\prime}}=1-\frac{\sum_{i=1}^{k} \sigma_{i}^{2}\left(1-\alpha_{i}\right)}{\sigma_{x}^{2}},
$$

where $\alpha_{i}$ is the reliability of the $i$ th strata, $\sigma_{i}^{2}$ is the variance between items in the $i$ th strata, and $\sigma_{x}^{2}$ is the variance of the total test scores. The stratified Cronbach alpha coefficient takes into account the weights proportional to the number of items and mean scores for each stratum. Qualls (1995) incorporated Raju's (1977) and Feldt's (Feldt \& Brennan, 1989) techniques for calculating reliability, which is called the Feldt-Raju coefficient.

The Feldt-Raju coefficient is defined as

$$
\text { Feldt-Raju } \rho_{X X^{\prime}}=\frac{\sigma_{x}^{2}-\sum_{i=1}^{k} \sigma_{i}^{2}}{\left(1-\sum_{i=1}^{k} \hat{\lambda}_{i}^{2}\right) \sigma_{x}^{2}},
$$

where $\sigma_{x}^{2}$ is the total score variance (i.e., the variance of the whole test); $\sigma_{i}^{2}$ indicates the score variance for a part-test (or item type) $i$; and $\hat{\lambda}_{i}$ is the sum of the variance of item type $i$ and the covariance between item type $i$ and other item types. This is defined as

$$
\hat{\lambda}_{i}=\frac{\left(\sigma_{i 1}+\sigma_{i 2}+\sigma_{i}^{2}+\sigma_{i(i+1)}+\cdots+\sigma_{i k}\right)}{\sigma_{x}^{2}}
$$

Table 2 through Table 3 display item types and their descriptions, as well as the number of items belonging to each item type. These tables were used to classify strata of item types. Because there were not large numbers of each of the individual item types, we organized the items into two categories for our analyses: multiple-choice and non-multiple choice.

Table 2: Reading Item Types and Descriptions

| Response Type | Description |
| :--- | :--- |
| multiplechoice (MC) | Student selects one correct answer from a number of options. |
| multipleselect (MS) | Student selects all correct answers from a number of options. |
| tablematch (MI) | Student checks a box to indicate if information from a column header matches <br> information from a row. |
| edittaskwithchoice (ETC) | Student identifies an incorrect word or phrase and chooses the replacement from a <br> number of options. |
| hottext (HT) | Student is directed to either select or use drag-and-drop feature to use text to support <br> an analysis or make an inference. |
| multiplechoice, <br> hottextselectable <br> (Two-part HT) | Student selects the correct answers from Part A and Part B. Part A is a multiple choice <br> or a multiselect, and Part B is a selectable HT. |
| Evidence-Based Selected | Student selects the correct answers from Part A and Part B. Part A often asks the <br> Response (EBSR) <br> to support Part A. |

Table 3: Mathematics Item Types and Descriptions

| Response Type | Description |
| :--- | :--- |
| multiplechoice (MC) | Student selects one correct answer from a number of options. |
| multipleselect (MS) | Student selects all correct answers from a number of options. |
| edittaskchoice (ETC) | Student identifies an incorrect word, phrase, or blank and chooses the replacement <br> from a number of options. |
| grid (GI) | The student selects numbers, words, phrases, or images and uses the drag-and-drop <br> feature to place them into a graphic. This item type may also require the student to <br> use the point, line, or arrow tools to create a response on a graph. |
| hottext (HT) | Student is directed to either select or use the drag-and-drop feature to use text to <br> support an analysis or make an inference. |
| equation (EQ) | Student uses a keypad with a variety of mathematical symbols to create a response. |
| textentrynaturallanguage | Student uses the keyboard to enter a response into a text field. |
| (NL) | Student checks a box to indicate if information from a column header matches <br> information from a row. |
| tablematch (MI) | Student types numeric values into a given table. |
| tableinput (TI) | An item that contains more than one response types. It could contain more than one <br> of the same response types or a combination of response types. |
| Multi-Interaction <br> (MULTI) |  |

Table 4: Reading Operational Item Types by Grade

| Item Type * |  |  | Grade |  |  |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :---: | :---: |
|  | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |  |  |
| MC | 28 | 29 | 21 | 24 | 31 | 29 | 29 | 35 |  |  |
| MS | 3 | 1 | 5 | 4 | 4 | 3 | 6 | 4 |  |  |
| ETC | 11 | 11 | 10 | 12 | 8 | 9 | 8 | 7 |  |  |
| HT | 4 | 1 | 5 | 6 | 5 | 7 | 3 | 2 |  |  |
| EBSR | 4 | 8 | 9 | 6 | 4 | 4 | 8 | 6 |  |  |

* Descriptions for each item type are presented in Table 2.

Table 5: Mathematics Operational Item Types by Grade

| Item Type * | Grade |  |  |  |  |  | Algebra 1** | Geometry** |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 3 | 4 | 5 | 6 | 7 | 8 |  |  |
| MC4 | 20 | 11 | 13 | 20 | 13 | 19 | 33; 29; 29; 28 | 22; 21; 18 |
| MS | 6 | 7 | 6 | 3 | 1 | 8 | 2; 1; 2; 3 | 5; 4; 7 |
| GI | 3 | 5 | 5 | 12 | 6 | 7 | 4; 5; 4; 5 | 2; 5; 3 |
| HT |  |  |  |  |  | 1 | 1; 0; 0; 0 | 5; 3; 2 |
| TI | 2 | 2 | 1 | 1 | 2 | 3 | 2; 3; 1; 1 | 0; 0; 0 |
| MI | 2 | 5 |  | 1 | 3 | 2 | 0; 1; 0; 1 | 0; 0; 0 |
| NL |  |  |  |  | 1 |  |  | 1; 1; 1 |
| EQ | 21 | 24 | 29 | 19 | 30 | 16 | 16; 19; 22; 20 | 21; 22; 25 |
| ETC |  |  |  |  |  |  |  | 2; 2; 2 |

* Descriptions for each item type are presented in Table 3.
** Algebra 1 has four core forms, and Geometry has three core forms.
Table 6 through Table 8 present the Cronbach alpha, stratified alpha, and Feldt-Raju coefficients for ELA, Mathematics, and EOC by grade/course and test form. The Cronbach alpha ranged from 0.88 to 0.92 for ELA, 0.90 to 0.95 for Mathematics, and 0.87 to 0.95 for EOC. The stratified alpha coefficients ranged from 0.88 to 0.92 for ELA, 0.91 to 0.95 for Mathematics, and 0.87 to 0.95 for EOC. The Feldt-Raju coefficients were between 0.85 and 0.91 for ELA, 0.89 and 0.94 for Mathematics, and 0.88 and 0.93 for EOC. The reliability coefficients by each demographic subgroup are presented in Appendix A. Reliability coefficients for each reporting category are also presented in Appendix A.

Table 6: Reliability Coefficients (ELA)

| Grade |  | Form | Cronbach Alpha | Stratified Alpha |
| :--- | :--- | :--- | :--- | :--- |
| 4 | Paper | 0.91 | 0.91 | 0.89 |
|  | Online | 0.90 | 0.90 | 0.88 |
|  | Accommodated | 0.88 | 0.88 | 0.85 |
| 6 | Online | 0.92 | 0.92 | 0.90 |
|  | Accommodated | 0.90 | 0.90 | 0.88 |
| 7 | Online | 0.91 | 0.91 | 0.90 |
|  | Accommodated | 0.89 | 0.89 | 0.86 |
| 8 | Online | 0.90 | 0.90 | 0.89 |
|  | Accommodated | 0.90 | 0.90 | 0.89 |
| 9 | Online | 0.92 | 0.92 | 0.90 |
|  | Accommodated | 0.91 | 0.91 | 0.88 |
| 10 | Online | 0.92 | 0.92 | 0.89 |
|  | Accommodated | 0.91 | 0.91 | 0.91 |

Table 7: Reliability Coefficients (Mathematics)

| Grade |  | Form | Cronbach Alpha | Stratified Alpha |
| :--- | :--- | :--- | :--- | :--- |
| 3 | Online | 0.94 | 0.94 | 0.93 |
|  | Accommodated | 0.94 | 0.94 | 0.92 |
| 4 | Online | 0.95 | 0.95 | 0.94 |
|  | Accommodated | 0.95 | 0.95 | 0.93 |
| 5 | Online | 0.95 | 0.95 | 0.93 |
|  | Accommodated | 0.94 | 0.94 | 0.90 |
| 6 | Online | 0.95 | 0.95 | 0.94 |
|  | Accommodated | 0.93 | 0.93 | 0.92 |
| 7 | Online | 0.94 | 0.94 | 0.92 |
|  | Accommodated | 0.93 | 0.93 | 0.89 |
| 8 | Online | 0.91 | 0.91 | 0.89 |
|  | Accommodated | 0.90 | 0.91 |  |

Table 8: Reliability Coefficients (EOC)

| Course | Form $^{*}$ | Cronbach Alpha | Stratified Alpha | Feldt-Raju |
| :---: | :--- | :--- | :--- | :--- |
| Algebra 1 | Online - Core 12 | 0.93 | 0.93 | 0.93 |
|  | Online - Core 13 | 0.93 | 0.93 | 0.93 |
|  | Online - Core 14 | 0.93 | 0.93 | 0.92 |


| Geometry | Online - Core 15 | 0.92 | 0.92 | 0.92 |
| :--- | :--- | :--- | :--- | :--- |
|  | Accommodated | 0.87 | 0.87 | 0.88 |
|  | Online - Core 9 | 0.94 | 0.94 | 0.93 |
|  | Online - Core 10 | 0.94 | 0.94 | 0.92 |
|  | Online - Core 11 | 0.95 | 0.95 | 0.91 |
|  | Accommodated | 0.88 | 0.88 | 0.89 |

* Each year, 3-4 core forms for EOC have been developed since Spring 2015. Each core form is assigned to a unique number.


### 3.2 Marginal Reliability

Marginal reliability is a measure of the overall reliability of the test based on the average conditional standard errors, estimated at different points on the achievement scale, for all students. The marginal reliability coefficients are nearly identical or close to coefficient alpha. For our analysis, the marginal reliability coefficients were computed using operational items.

Within the IRT framework, measurement error varies across the range of ability. The amount of precision is indicated by the test information at any given point of a distribution. The inverse of the test information function represents the standard error of measurement. The standard error of measurement is equal to the inverse square root of information. The larger the measurement error, the less test information is being provided. The amount of test information provided is at its maximum for students toward the center of the distribution, as opposed to students with more extreme scores. Conversely, measurement error is minimal for the part of the underlying scale that is at the middle of the test distribution and greater on scaled values farther away from the middle.

The marginal reliability is defined as:

$$
\bar{\rho}=1-\frac{\int \sigma_{e}^{2}(\widehat{\theta}) f(\hat{\theta}) d \hat{\theta}}{\sigma_{x}^{2}}
$$

where $\sigma_{e}^{2}(\hat{\theta})$ is the function generating the standard error of measurement and $f(\hat{\theta})$ is the assumed population density. The marginal reliability of a test is computed by integrating $\theta$ out of the test information function as follows:

$$
\rho=\frac{\sigma_{\theta}^{2}-\bar{\sigma}_{e}^{2}}{\sigma_{\theta}^{2}}
$$

where $\sigma_{\theta}^{2}$ is the true score variance of $\theta$ and

$$
\bar{\sigma}_{e}^{2}=\int_{-\infty}^{\infty} \frac{1}{I(\theta)} g(\theta) d \theta
$$

where $g(\theta)$ is a density function. If population parameters are assumed normal, then $g(\theta) \sim N\left(\mu, \sigma^{2}\right)$. In the absence of information about the population distribution of $\theta$, a uniform prior is available such that $g(\theta) \sim \mathrm{U}[\mathrm{a}, \mathrm{b}]$ where a and b are the lower and upper limits of the uniform distribution, respectively. The integral is evaluated using Gauss-Hermite quadrature:

$$
\bar{\sigma}_{e}^{2} \approx \sum_{q=1}^{Q} \frac{1}{I\left(\theta_{q}\right)} w_{q}
$$

where $\theta_{q}$ is the value at node $q$ and $w_{q}$ is the weight at node $q$. The true score variance of $\theta$ can be obtained from the marginal maximum likelihood (MML) means procedure.

In item response theory (IRT), the marginal likelihood is typically maximized to estimate item parameters by integrating $\theta$ out of the function and treating population parameters as known. However, suppose the item parameters are treated as fixed but the population parameters are treated as latent. Then, the following marginal likelihood can be maximized with respect to the two latent parameters associated with the normal population distribution:

$$
\arg \max L(\mu, \sigma)=\prod_{i=1}^{N} \int_{-\infty}^{\infty} \prod_{j=1}^{K} p\left(x_{j} \mid \theta_{i}, \mathbf{Y}_{\boldsymbol{j}}\right) g(\theta \mid \mu, \sigma) d \theta
$$

where in this context $p\left(x_{j} \mid \theta_{i}, \mathbf{Y}_{\boldsymbol{j}}\right)$ is used to mean the probability of individual $i=\{1,2, \ldots, N\}$ having observed response $x$ to item $j=\{1,2, \ldots, K\}$ given the vector of item parameters, $\boldsymbol{\Upsilon}$. The integral has no closed form and so the function is evaluated using a fixed quadrature routine. Rather than using Gauss-Hermite, $Q$ nodes are chosen from the normal distribution at fixed points and then the integral is evaluated by summation over the $Q$ nodes as:

$$
\arg \max L(\mu, \sigma)=\prod_{i=1}^{N} \sum_{q=1}^{Q} \prod_{j=1}^{K} p\left(x_{j} \mid \theta_{q}, \mathbf{r}_{j}\right) g\left(\theta_{q} \mid \mu, \sigma\right)
$$

where $\theta_{q}$ is node q . In this instance, fixed quadrature points allow a smaller number of likelihood evaluations because the values for $\theta_{q}$ are fixed. If Gauss-Hermite were used, the nodes would change as each value of $\mu$ and $\sigma$ are updated and the likelihood calculations would need to be performed at each iteration.

Table 9 presents the marginal reliability coefficients for all students. The marginal reliability coefficients for all subjects and grades were ranging from 0.85 to 0.93 .

Table 9: Marginal Reliability Coefficients

| Subject | Grade | Marginal Reliability for Response Pattern Scores | Subject | Grade/Course | Marginal Reliability for Response Pattern Scores |
| :---: | :---: | :---: | :---: | :---: | :---: |
| ELA | 3 | 0.90 | Mathematics | 3 | 0.93 |
|  | 4 | 0.88 |  | 4 | 0.92 |
|  | 5 | 0.89 |  | 5 | 0.93 |
|  | 6 | 0.88 |  | 6 | 0.93 |
|  | 7 | 0.85 |  | 7 | 0.92 |
|  | 8 | 0.88 |  | 8 | 0.89 |
|  | 9 | 0.88 | EOC | Algebra 1 | 0.93 |
|  | 10 | 0.88 |  | Geometry | 0.93 |

### 3.3 Test Information Curves and Standard Error of Measurement

Within the IRT framework, measurement error varies across the range of ability as a result of the test information function (TIF). The TIF describes the amount of information provided by the test at each score point along the ability continuum. The inverse of the TIF is characterized as the conditional measurement error at each score point. For instance, if the measurement error is large, then less information is being provided by the assessment at the specific ability level.

Figure 1 displays a sample TIF from the FSA. The graphic shows that this test information is maximized in the middle of the score distribution, meaning it provides the most precise scores in this range. Where the curve is lower at the tails indicates that the test provides less information about examinees at the tails relative to the center. The vertical lines are samples of the performance cuts.

Figure 1: Sample Test Information Function


Computing these TIFs is useful to evaluate where the test is maximally informative. In IRT, the TIF is based on the estimates of the item parameters in the test, and the formula used for the FSA is calculated as:

$$
\begin{aligned}
\operatorname{TIF}\left(\theta_{i}\right)= & \sum_{j=1}^{N_{G P C M}} D^{2} a_{j}^{2}\left(\frac{\sum_{s=1}^{m_{j}} s^{2} \operatorname{Exp}\left(\sum_{h=1}^{s} D a_{j}\left(\theta_{i}-b_{j h}\right)\right)}{1+\sum_{s=1}^{m_{j}} \operatorname{Exp}\left(\sum_{h=1}^{s} D a_{j}\left(\theta_{i}-b_{j h}\right)\right)}\right. \\
& \left.\quad-\left(\frac{\sum_{s=1}^{m_{j}} \operatorname{sexp}\left(\sum_{h=1}^{s} D a_{j}\left(\theta_{i}-b_{j h}\right)\right)}{1+\sum_{s=1}^{m_{j}} \operatorname{Exp}\left(\sum_{h=1}^{s} D a_{j}\left(\theta_{i}-b_{j h}\right)\right)}\right)^{2}\right)+\sum_{j=1}^{N_{3 P L}} D^{2} a_{j}^{2}\left(\frac{Q_{j}}{P_{j}}\left[\frac{P_{j}-c_{j}}{1-c_{j}}\right]^{2}\right),
\end{aligned}
$$

where $N_{G P C M}$ is the number of items that are scored using generalized partial credit model (GPCM) items, $N_{3 P L}$ is the number of items scored using 3PL or 2 PL model, $j$ indicates item $j(j \in$ $\{1,2, \ldots, N\}), m_{j}$ is the maximum possible score of the item, $s$ indexes step of the item, $b_{j h}$ is the $h$ th step for item $j$ with $m$ total categories. $\theta_{i}$ is the ability of student $i$.

The standard error for estimated student ability (theta score) is the square root of the reciprocal of the TIF:

$$
\operatorname{se}\left(\theta_{i}\right)=\frac{1}{\sqrt{\operatorname{TIF}\left(\theta_{i}\right)}}
$$

It is typically more useful to consider the inverse of the TIF rather than the TIF itself, as the standard errors are more useful for score interpretation. For this reason, standard error plots are presented in Figure 2, Figure 3, and Figure 4, respectively, instead of the TIFs for ELA, Mathematics, and EOC. These plots are based on the scaled scores reported in 2018. Vertical lines represent the four performance category cut scores.

Figure 2: Conditional Standard Errors of Measurement (ELA)


Grade4




Figure 3: Conditional Standard Errors of Measurement (Mathematics) Grade3

Grade4




Figure 4: Conditional Standard Errors of Measurement (EOC)

Algebra1


Geometry


For most tests, the standard error curves follow the typical expected trends with more test information regarding scores observed near the middle of the score scale. However, there are two general exceptions. In grade 8 Mathematics and for both Algebra EOC tests, the test is maximized at a higher point along the ability scale. This suggests the items comprising these tests are somewhat challenging relative to the tested population.

Appendix B includes scale score by scale score conditional standard errors of measurement and corresponding achievement levels for each scale score.

In classical test theory, the SEM is defined as $s_{x} \sqrt{1-r_{x x \prime}}$, where $s_{x}$ is the standard deviation of the raw score, and $r_{x x}$, is the reliability coefficient. Under classical test theory, measurement error is assumed to be the same at all levels of achievement, and one reliability coefficient can be estimated to acknowledge that error. Standard error of measurement indicates the standard deviation of a single student's repeated test scores, if he or she were to take the same test repeatedly (with no new learning or no memory of questions taking place between test administrations). Reliability coefficients and SEM for each reporting category are also presented in Appendix A.

### 3.4 Reliability of Achievement Classification

When students complete the FSA, they are placed into one of five achievement levels given their observed scaled score. The cut scores for student classification into the different achievement levels were determined after the FSA standard-setting process.

During test construction, techniques are implemented to minimize misclassification of students, which can occur on any assessment. In particular, standard error of measurement (SEM) curves can be constructed to ensure that smaller SEMs are expected near important cut scores of the test.

### 3.4.1 Classification Accuracy

Misclassification probabilities are computed for all achievement level standards (i.e., for the cuts between levels 1 and 2, levels 2 and 3, levels 3 and 4, and levels 4 and 5). The achievement level cut between level 2 and level 3 is of primary interest because students are classified as Satisfactory or Below Satisfactory using this cut. Students with observed scores far from the level 3 cut are expected to be classified more accurately as Satisfactory or Below Satisfactory than students with scores near this cut. This report estimates classification reliabilities using two different methods: one based on observed abilities and a second based on estimating a latent posterior distribution for the true scores.

Two approaches for estimating classification probabilities are provided. The first is an observed score approach to computing misclassification probabilities and is designed to explore the following research questions:

1. What is the overall classification accuracy index (CAI) of the total test?
2. What is the classification accuracy rate index for each individual performance cut within the test?

The second approach computes misclassification probabilities using an IRT-based method for students scoring at each score point. This approach is designed to explore the following research questions:

1. What is the probability that the student's true score is below the cut point?
2. What is the probability that the student's true score is above the cut point?

Both approaches yield student-specific classification probabilities that can be aggregated to form overall misclassification rates for the test.

In the first approach, we used students from the spring 2018 FSA population data files with the status of reported scores. However, in the second approach, item level data from the calibration sample were used. Since there were multiple core forms in EOC tests, the classification accuracy analysis was performed for each form, as operational items varied by form. Also, the item level data used in the IRT-based approach did not include accommodated tests because the sample was too small to compute classification accuracy.

Table 10 provides the sample size, mean, and standard deviation of the observed theta for the data used in the first method described earlier. The theta scores are based on the maximum likelihood estimates (MLEs) obtained from AIR's scoring engine. Similarly, Table 11 provides the sample size, mean, and standard deviation of the observed theta for the data used in the second method.

Table 10: Descriptive Statistics from Population Data

| ELA Grade | Sample <br> Size | Average <br> Theta | Standard <br> Deviation of <br> Theta | Mathematics <br> Grade/EOC <br> Subject | Sample <br> Size | Average <br> Theta | Standard <br> Deviation of <br> Theta |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 3 | 221466 | 0.09 | 1.07 | 3 | 220988 | 0.08 | 1.07 |
| 4 | 213887 | 0.04 | 1.02 | 4 | 215313 | 0.06 | 1.12 |
| 5 | 209361 | 0.05 | 1.09 | 5 | 210682 | 0.12 | 1.10 |
| 6 | 210089 | -0.01 | 1.13 | 6 | 201233 | -0.08 | 1.12 |
| 7 | 200194 | -0.03 | 1.12 | 7 | 169887 | -0.06 | 1.19 |
| 8 | 201003 | 0.05 | 1.11 | 8 | 116606 | -0.18 | 1.16 |
| 9 | 198261 | 0.01 | 1.09 | Algebra 1 | 204200 | 0.10 | 1.19 |
| 10 | 192049 | 0.05 | 1.11 | Geometry | 188117 | 0.04 | 1.13 |

Table 11: Descriptive Statistics from Calibration Data

| ELA |  |  |  | Mathematics |  |  |  | EOC |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Grade | N | Average <br> Theta | SD of <br> Theta | Grade | $N$ | Average <br> Theta | SD of <br> Theta | Subject/Core | $N$ | Average <br> Theta | SD of <br> Theta |
| 3 | 26454 | 0.09 | 1.08 | 3 | 25488 | 0.08 | 1.06 | Alg1/Core12 | 40360 | -0.04 | 1.21 |
| 4 | 23490 | 0.02 | 1.05 | 4 | 25920 | 0.09 | 1.12 | Alg1/Core13 | 32685 | 0.05 | 1.16 |
| 5 | 22097 | 0.06 | 1.09 | 5 | 26088 | 0.15 | 1.1 | Alg1/Core14 | 32522 | 0.09 | 1.15 |
| 6 | 23046 | 0.02 | 1.14 | 6 | 26777 | 0.00 | 1.11 | Alg1/Core15 | 27853 | 0.06 | 1.16 |
| 7 | 21178 | -0.02 | 1.12 | 7 | 23456 | -0.04 | 1.18 | Geo/Core9 | 55763 | -0.05 | 1.11 |
| 8 | 24716 | 0.09 | 1.09 | 8 | 93872 | -0.17 | 1.15 | Geo/Core10 | 40215 | 0.02 | 1.08 |
| 9 | 24778 | 0.06 | 1.07 |  |  |  |  | Geo/Core11 | 40338 | 0.03 | 1.06 |
| 10 | 23823 | 0.08 | 1.11 |  |  |  |  |  |  |  |  |

The observed score approach (Rudner, 2001) implemented to assess classification accuracy is based on the probability that the true score, $\theta_{\text {, for student } i}$ is within performance level $j=$ $1,2, \cdots, J$. This probability can be estimated from evaluating the following integral

$$
p_{i j}=\operatorname{Pr}\left(\lambda_{l} \leq \theta_{i}<\lambda_{u} \mid \hat{\theta}_{i}, \hat{\sigma}_{i}^{2}\right)=\int_{\lambda_{l}}^{\lambda_{u}} f\left(\theta_{i} \mid \hat{\theta}_{i}, \hat{\sigma}_{i}^{2}\right) d \theta_{i,}
$$

where $\lambda_{u}$ and $\lambda_{l}$ denote the score corresponding to the upper and lower limits of the performance level, respectively, $\widehat{\theta}_{i}$ is the ability estimate of the $i$ th student with standard error of measurement of $\hat{\sigma}_{i}$ and using the asymptotic property of normality of the maximum likelihood estimate, $\hat{\theta}_{i}$, we take $f(\cdot)$ as asymmetrically normal, so the above probability can be estimated by

$$
p_{i j}=\Phi\left(\frac{\lambda_{u}-\hat{\theta}_{i}}{\hat{\sigma}_{i}}\right)-\Phi\left(\frac{\lambda_{l}-\hat{\theta}_{i}}{\hat{\sigma}_{i}}\right)
$$

where $\Phi(\cdot)$ denotes the standard normal cumulative distribution function (CDF).
The expected number of students at level $j$ based on students from observed level $k$ can be expressed as

$$
E_{k j}=\sum_{p l_{i} \epsilon k} p_{i j}
$$

where $p l_{i}$ is the $i$ th student's performance level, the values of $E_{k j}$ are the elements used to populate the matrix $\boldsymbol{E}$, a $5 \times 5$ matrix of conditionally expected numbers of students to score within each performance level bin based on their true scores. The overall CAI of the test can then be estimated from the diagonal elements of the matrix:

$$
\mathrm{CAI}=\frac{\operatorname{tr}(\boldsymbol{E})}{N}
$$

where $N=\sum_{k=1}^{5} N_{k}, N_{k}$ is the observed number of students scoring in performance level $k$. The classification accuracy index for the individual cuts (CAIC) is estimated by forming square partitioned blocks of the matrix $\boldsymbol{E}$ and taking the summation over all elements within the block as follows:

$$
\mathrm{CAIC}=\left(\sum_{k=1}^{p} \sum_{j=1}^{p} E_{k j}+\sum_{k=p+1}^{5} \sum_{j=p+1}^{5} E_{k j}\right) / N
$$

where $p$ is the element of one of the cuts of interest.
The IRT-based approach (Guo, 2006) makes use of student-level item response data from the 2018 FSA test administration. We can estimate a posterior probability distribution for the latent true score and from this estimate the probability that a true score is above the cut as

$$
p(\theta>c)=\frac{\int_{c}^{\infty} p(z \mid \theta) f(\theta \mid \mu, \sigma) d \theta}{\int_{-\infty}^{\infty} p(z \mid \theta) f(\theta \mid \mu, \sigma) d \theta}
$$

where $c$ is the cut score required for passing in the same assigned metric, $\theta$ is true ability in the true-score metric, $z$ is the item score, $\mu$ is the mean, and $\sigma$ is the standard deviation of the population distribution. The function $p(z \mid \theta)$ is the probability of the particular pattern of responses given the theta, and $f(\theta)$ is the density of the proficiency $\theta$ in the population.

Similarly, we can estimate the probability that a true score is below the cut as

$$
p(\theta<c)=\frac{\int_{-\infty}^{c} p(z \mid \theta) f(\theta \mid \mu, \sigma) d \theta}{\int_{-\infty}^{\infty} p(z \mid \theta) f(\theta \mid \mu, \sigma) d \theta}
$$

From these misclassification probabilities, we can estimate the overall false positive rate (FPR) and false negative rate (FNR) of the test. The FPR is expressed as the proportion of individuals who scored above the cut based on their observed score, but their true score would otherwise have classified them as below the cut. The FNR is expressed as the proportion of individuals who scored below the cut based on their observed score, but otherwise would have been classified as above the cut based on their true scores. These rates are estimated as follows:

$$
\mathrm{FPR}=\sum_{i \in \theta \geq c} p(\theta<c) / N
$$

$$
\mathrm{FNR}=\sum_{i \in \theta<c} p(\theta \geq c) / N
$$

In addition to these rates, we computed the accuracy rates for each cut as

$$
\text { Accuracy = } 1 \text { - (FPR + FNR). }
$$

Table 12 and Table 13 provide the overall classification accuracy index (CAI) and the classification accuracy index for the individual cuts (CAIC) for the ELA and Mathematics tests, respectively, based on the observed score approach. Here, the overall classification accuracy of the test ranges from 0.763 to around 0.804 for Mathematics and EOC, and from 0.731 to 0.754 for ELA.

The overall cut accuracy rates are much higher, denoting that the degree to which we can reliably differentiate students between adjacent performance levels is typically above or close to 0.9 .

Table 12: Classification Accuracy Index (ELA)

| Grade | Overall Accuracy <br> Index | Cut Accuracy Index |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
|  |  | Cut 2 and Cut 3 | Cut 3 and Cut 4 | Cut 4 and Cut 5 |  |
| 3 | 0.744 | 0.943 | 0.920 | 0.922 | 0.957 |
| 4 | 0.737 | 0.933 | 0.916 | 0.925 | 0.961 |
| 5 | 0.748 | 0.946 | 0.923 | 0.923 | 0.955 |
| 6 | 0.738 | 0.935 | 0.919 | 0.925 | 0.957 |
| 7 | 0.731 | 0.928 | 0.917 | 0.928 | 0.956 |
| 8 | 0.746 | 0.948 | 0.927 | 0.924 | 0.945 |
| 9 | 0.753 | 0.943 | 0.924 | 0.927 | 0.957 |
| 10 | 0.754 | 0.947 | 0.924 | 0.926 | 0.956 |

Table 13: Classification Accuracy Index (Mathematics and EOC)

| Grade/Subject | Overall Accuracy <br> Index | Cut Accuracy Index |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
|  |  | Cut 2 and Cut 3 | Cut 3 and Cut 4 | Cut 4 and Cut 5 |  |
| 3 | 0.776 | 0.959 | 0.936 | 0.926 | 0.954 |
| 4 | 0.776 | 0.954 | 0.940 | 0.932 | 0.948 |
| 5 | 0.799 | 0.957 | 0.944 | 0.941 | 0.957 |
| 6 | 0.804 | 0.947 | 0.936 | 0.949 | 0.973 |
| 7 | 0.801 | 0.939 | 0.940 | 0.951 | 0.970 |
| 8 | 0.773 | 0.916 | 0.919 | 0.957 | 0.981 |
| Algebra 1 | 0.763 | 0.926 | 0.923 | 0.938 | 0.963 |
| Geometry | 0.776 | 0.919 | 0.921 | 0.955 | 0.973 |

Table 14, Table 15, and Table 16 provide the FPR and FNR from the IRT-based approach for both ELA and Mathematics tests. In ELA, the FPR and FNR rates for the level $2 / 3$ cut are around $7 \%$ to $9 \%$. In Mathematics, the FNR rates for the level $2 / 3$ cut are around $5 \%$ to $6 \%$ with the exception of grade 8 Mathematics, and the FPR rates are around $6 \%$ to $8 \%$. In EOC, the rates are around $7 \%$ to $10 \%$. Table 14 , Table 15 , and Table 16 also provide the overall accuracy rates after accounting for both false positive and false negative rates. For example, the overall accuracy rate of 0.866 for the level $2 / 3$ cut in grade 3 Mathematics suggests $86.6 \%$ of the students estimated to have a true score status at level 3 are correctly classified into that category by their observed scores. As expected, the overall accuracy rates are reasonable in all cuts except at the extreme cuts. A high false negative rate at the cut between levels 4 and 5 is also expected due to large standard error.

Table 14: False Classification Rates and Overall Accuracy Rates (ELA)

|  | 1/2 cut |  |  | 2/3 cut |  |  | 3/4 cut |  |  | $4 / 5$ cut |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Grade | FPR | FNR | Accuracy | FPR | FNR | Accuracy | FPR | FNR | Accuracy | FPR | FNR | Accuracy |
| 3 | 0.151 | 0.034 | 0.815 | 0.082 | 0.077 | 0.841 | 0.045 | 0.169 | 0.787 | 0.017 | 0.343 | 0.640 |
| 4 | 0.161 | 0.041 | 0.798 | 0.085 | 0.081 | 0.834 | 0.042 | 0.156 | 0.802 | 0.017 | 0.303 | 0.681 |
| 5 | 0.147 | 0.031 | 0.822 | 0.081 | 0.076 | 0.843 | 0.046 | 0.166 | 0.788 | 0.018 | 0.327 | 0.655 |
| 6 | 0.152 | 0.038 | 0.810 | 0.081 | 0.087 | 0.833 | 0.046 | 0.153 | 0.801 | 0.017 | 0.283 | 0.699 |
| 7 | 0.154 | 0.047 | 0.800 | 0.081 | 0.088 | 0.830 | 0.043 | 0.147 | 0.810 | 0.019 | 0.251 | 0.730 |
| 8 | 0.134 | 0.032 | 0.834 | 0.082 | 0.066 | 0.851 | 0.049 | 0.145 | 0.806 | 0.024 | 0.279 | 0.698 |
| 9 | 0.128 | 0.036 | 0.836 | 0.077 | 0.076 | 0.847 | 0.047 | 0.134 | 0.819 | 0.020 | 0.250 | 0.730 |
| 10 | 0.122 | 0.033 | 0.845 | 0.076 | 0.078 | 0.846 | 0.045 | 0.141 | 0.814 | 0.020 | 0.277 | 0.703 |

Table 15: False Classification Rates and Overall Accuracy Rates (Mathematics)

|  | $\mathbf{1 / 2}$ cut |  |  | 2/3 cut |  |  | 3/4 cut |  |  | 4/5 cut |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Grade | FPR | FNR | Accuracy | FPR | FNR | Accuracy | FPR | FNR | Accuracy | FPR | FNR | Accuracy |
| 3 | 0.104 | 0.025 | 0.872 | 0.080 | 0.054 | 0.866 | 0.050 | 0.128 | 0.822 | 0.019 | 0.293 |  |
| 4 | 0.112 | 0.027 | 0.861 | 0.075 | 0.048 | 0.877 | 0.045 | 0.107 | 0.847 | 0.025 | 0.259 | 0.717 |
| 5 | 0.110 | 0.024 | 0.865 | 0.070 | 0.047 | 0.883 | 0.042 | 0.090 | 0.868 | 0.020 | 0.193 | 0.786 |
| 6 | 0.112 | 0.030 | 0.859 | 0.065 | 0.062 | 0.872 | 0.034 | 0.101 | 0.864 | 0.013 | 0.202 | 0.785 |
| 7 | 0.124 | 0.036 | 0.840 | 0.059 | 0.058 | 0.884 | 0.028 | 0.111 | 0.861 | 0.013 | 0.211 | 0.776 |
| 8 | 0.144 | 0.061 | 0.795 | 0.064 | 0.110 | 0.826 | 0.020 | 0.168 | 0.811 | 0.008 | 0.227 | 0.765 |

Table 16: False Classification Rates and Overall Accuracy Rates (EOC)

|  | 1/2 cut |  |  | 2/3 cut |  |  | 3/4 cut |  |  | 4/5 cut |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Grade | FPR | FNR | Accuracy | FPR | FNR | Accuracy | FPR | FNR | Accuracy | FPR | FNR | Accuracy |
| Algebra 1 Core12 | 0.147 | 0.054 | 0.799 | 0.088 | 0.072 | 0.840 | 0.033 | 0.112 | 0.856 | 0.014 | 0.146 | 0.840 |
| Algebra 1 Core13 | 0.154 | 0.050 | 0.796 | 0.094 | 0.072 | 0.834 | 0.036 | 0.114 | 0.850 | 0.015 | 0.145 | 0.840 |
| Algebra 1 Core14 | 0.139 | 0.048 | 0.813 | 0.092 | 0.065 | 0.842 | 0.036 | 0.110 | 0.854 | 0.017 | 0.155 | 0.828 |
| Algebra 1 Core15 | 0.157 | 0.054 | 0.789 | 0.101 | 0.074 | 0.825 | 0.036 | 0.119 | 0.845 | 0.015 | 0.148 | 0.837 |
| Geometry Core9 | 0.170 | 0.059 | 0.771 | 0.086 | 0.080 | 0.834 | 0.023 | 0.108 | 0.869 | 0.011 | 0.139 | 0.850 |
| Geometry Core10 | 0.163 | 0.052 | 0.784 | 0.085 | 0.073 | 0.841 | 0.024 | 0.101 | 0.875 | 0.012 | 0.135 | 0.852 |
| Geometry Core11 | 0.156 | 0.051 | 0.793 | 0.088 | 0.074 | 0.838 | 0.024 | 0.108 | 0.868 | 0.012 | 0.141 | 0.847 |

Figure 5 shows a plot exhibiting the probability of misclassification for grade 3 ELA. The plot displays that students with scores below -0.334 on the theta scale, which corresponds to a scale score of 293 , and students with scores above 0.361 , corresponding to a scale score of 307 , are classified accurately at least $90 \%$ of the time. Scale scores representing $90 \%$ of classification accuracy by each grade and subject are displayed in Appendix D.

Appendix D also includes plots of the misclassification probabilities for the level $2 / 3$ cuts from the IRT-based approach conditional on ability for all grades and subject as well as by subgroups (English Language Learners [ELLs] and Students with Disabilities [SWDs]). The vertical bar within each graph represents the cut score required to achieve level 3 (i.e., satisfactory). A properly functioning test yields increased misclassification probabilities approaching the cut, as the density of the posterior probability distribution is symmetric, and approximately half of its mass will fall on either side of the proficiency level cut as $\theta \rightarrow c$.

These visual displays are useful heuristics to evaluate the probability of misclassification for all levels of ability. Students far from the level 3 cut have very small misclassification probabilities, and the probabilities approach a peak near $50 \%$ as $\theta \rightarrow c$, as expected.

Figure 5: Probability of Misclassification Conditional on Ability

## Probability of Misclassification Conditional on Ability



These results demonstrate that classification reliabilities are generally high, with some lower rates affecting tests known to be particularly challenging. The classification accuracy results presented in this report (Table 12 and Table 13) are generally equivalent to or higher than those reported in the 2013 FCAT 2.0 and EOC technical reports. Based on the Florida Statewide Assessments 2013 Yearbook (Florida Department of Education, 2013), the classification accuracy rates in Mathematics ranged from 0.690 in grade 4 to 0.719 in grade 5 (see page 112 for details). Similarly, the classification accuracy rates in Reading ranged from 0.664 in grade 10 to 0.718 in grade 3 (see page 264 for details). The classification accuracy rates in Algebra 1 vary from 0.716 to 0.737 (see page 413 for details). Additionally, we can compare the FSA classification accuracy rates to those of the State of New York, which is comparable in population size (New York State Education Department, 2014). Although New York administers a different testing program, estimated accuracy rates there range from $77 \%$ to $81 \%$ in ELA and from $81 \%$ to $85 \%$ in Mathematics (see page 100 for details). While the overall cut accuracy results for New York are slightly higher than those of the FSA, as there are only three achievement level cuts compared to four FSA cuts, the individual cut accuracy was comparable between New York and Florida. Florida showed from $92 \%$ to $93 \%$ in ELA and from $92 \%$ to $94 \%$ in Mathematics and EOC for the level $2 / 3$ cut. New York showed from $91 \%$ to $93 \%$ in ELA and from $93 \%$ to $95 \%$ in Mathematics and EOC for the proficiency cut.

### 3.4.2 Classification Consistency

Classification accuracy refers to the degree to which a student's true score and observed score would fall within the same performance level (Rudner, 2001). Classification consistency refers to the degree to which examinees are classified into the same performance level assuming the test is administered twice independently (Lee, Hanson, and Brennan, 2002) - that is, the percentages of students who are consistently classified in the same performance levels on two equivalent test forms. In reality, the true ability is unknown, and students do not take an alternate, equivalent form; therefore, classification accuracy and consistency are estimated based on students' item scores and the item parameters, and the assumed underlying latent ability distribution. Classification consistency was estimated based on the method in Lee, Hanson, and Brennan (2002).

Similar to accuracy, a $5 \times 5$ matrix can be constructed by assuming the test is administered twice independently to the same group of students. The classification consistency index for the individual cuts (CCIC) was estimated. Classification consistency with classification accuracy results are presented in Tables 17-20. In cut 1 and cut 2, cut 2 and cut 3 , and cut 3 and cut 4 results, all accuracy values are higher than .9 , and consistency values are around .9 or slightly below .9 . With the higher performance levels, cut 4 and cut 5, most values are higher than .95 or around .95 . In all performance levels, classification accuracy is slightly higher than classification consistency. Classification consistency rates can be lower than classification accuracy because the consistency is based on two tests with measurement errors, while the accuracy is based on one test with a measurement error and the true score. The accuracy and consistency rates for each performance level are higher for the levels with smaller standard error.

Table 17. Classification Accuracy and Consistency (Cut 1 and Cut 2)

| Grade | ELA |  | Grade/ <br> Subject | Mathematics |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Accuracy | Consistency |  | Accuracy | Consistency |
| 3 | 0.943 | 0.918 | 3 | 0.959 | 0.943 |
| 4 | 0.933 | 0.905 | 4 | 0.954 | 0.936 |
| 5 | 0.946 | 0.924 | 5 | 0.957 | 0.942 |
| 6 | 0.935 | 0.910 | 6 | 0.947 | 0.930 |
| 7 | 0.928 | 0.896 | 7 | 0.939 | 0.916 |
| 8 | 0.948 | 0.926 | 8 | 0.916 | 0.876 |
| 9 | 0.943 | 0.919 | Algebra 1 | 0.926 | 0.888 |
| 10 | 0.947 | 0.926 | Geometry | 0.919 | 0.881 |

Table 18. Classification Accuracy and Consistency (Cut 2 and Cut 3)

| Grade | ELA |  | Grade/ <br> Subject | Mathematics |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Accuracy | Consistency |  | Accuracy | Consistency |
| 3 | 0.920 | 0.888 | 3 | 0.936 | 0.910 |
| 4 | 0.916 | 0.883 | 4 | 0.940 | 0.918 |
| 5 | 0.923 | 0.889 | 5 | 0.944 | 0.921 |
| 6 | 0.919 | 0.882 | 6 | 0.936 | 0.910 |
| 7 | 0.917 | 0.881 | 7 | 0.940 | 0.918 |
| 8 | 0.927 | 0.896 | 8 | 0.919 | 0.885 |
| 9 | 0.924 | 0.892 | Algebra 1 | 0.923 | 0.887 |
| 10 | 0.924 | 0.891 | Geometry | 0.921 | 0.886 |

Table 19. Classification Accuracy and Consistency (Cut 3 and Cut 4)

| Grade | ELA |  | Grade/ | Mathematics |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
|  | Accuracy | Consistency |  | Accuracy | Consistency |
| 3 | 0.922 | 0.890 | 3 | 0.926 | 0.894 |
| 4 | 0.925 | 0.897 | 4 | 0.932 | 0.905 |
| 5 | 0.923 | 0.890 | 5 | 0.941 | 0.917 |
| 6 | 0.925 | 0.892 | 6 | 0.949 | 0.924 |
| 7 | 0.928 | 0.899 | 7 | 0.951 | 0.931 |
| 8 | 0.924 | 0.889 | 8 | 0.957 | 0.942 |
| 9 | 0.927 | 0.896 | Algebra 1 | 0.938 | 0.920 |
| 10 | 0.926 | 0.895 | Geometry | 0.955 | 0.941 |

Table 20. Classification Accuracy and Consistency (Cut 4 and Cut 5)

| Grade | ELA |  | Grade/ <br> Subject | Mathematics |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Accuracy | Consistency |  | Accuracy | Consistency |
| 3 | 0.957 | 0.943 | 3 | 0.954 | 0.939 |
| 4 | 0.961 | 0.949 | 4 | 0.948 | 0.926 |
| 5 | 0.955 | 0.943 | 5 | 0.957 | 0.939 |
| 6 | 0.957 | 0.943 | 6 | 0.973 | 0.960 |
| 7 | 0.956 | 0.940 | 7 | 0.970 | 0.958 |
| 8 | 0.945 | 0.925 | 8 | 0.981 | 0.976 |
| 9 | 0.957 | 0.940 | Algebra 1 | 0.963 | 0.955 |
| 10 | 0.956 | 0.940 | Geometry | 0.973 | 0.967 |

### 3.5 Precision at Cut Scores

Table 21 through Table 23 present mean conditional standard error of measurement at each achievement level by grade and subject. These tables also include achievement level cut scores and associated conditional standard error of measurement.

Table 21: Achievement Levels and Associated Conditional Standard Error of Measurement (ELA)

| Grade | Achievement Level | Mean CSEM | Cut Score (Scale Score) | CSEM at Cut Score |
| :---: | :---: | :---: | :---: | :---: |
| 3 | 1 | 8.24 |  |  |
| 3 | 2 | 5.54 | 285 | 6 |
| 3 | 3 | 5.24 | 300 | 5 |
| 3 | 4 | 6.17 | 315 | 6 |
| 3 | 5 | 8.38 | 330 | 7 |
| 4 | 1 | 8.63 |  |  |
| 4 | 2 | 5.86 | 297 | 7 |
| 4 | 3 | 5.16 | 311 | 6 |
| 4 | 4 | 5.25 | 325 | 6 |
| 4 | 5 | 6.46 | 340 | 6 |
| 5 | 1 | 7.66 |  |  |
| 5 | 2 | 5.86 | 304 | 7 |
| 5 | 3 | 5.93 | 321 | 6 |
| 5 | 4 | 6.45 | 336 | 7 |
| 5 | $5$ | 8.12 | 352 | 8 |
| 6 | 1 | 8.46 |  |  |
| 6 | 2 | 6.53 | 309 | 8 |
| 6 | 3 | 6.05 | 326 | 7 |
| 6 | 4 | 6.16 | 339 | 7 |
| 6 | $5$ | 7.52 | 356 | 7 |
| 7 | 1 | 9.35 |  |  |
| 7 | 2 | 6.53 | 318 | 8 |
| 7 | 3 | 5.99 | 333 | 7 |
| 7 | 4 | 5.97 | 346 | 6 |
| 7 | 5 | 6.89 | 360 | 7 |
| 8 | 1 | 7.52 |  |  |
| 8 | $2$ | 5.49 | 322 | 6 |
| 8 | 3 | 5.66 | 337 | 6 |
| 8 | $4$ | 6.29 | $352$ | 6 |
| 8 | 5 | 8.37 | 366 | 8 |
| 9 | 1 | 8.00 |  |  |
| 9 | 2 | 5.60 | 328 | 7 |
| 9 | 3 | 5.03 | 343 | 6 |
| 9 | 4 | 5.40 | 355 | 6 |


| Grade | Achievement <br> Level | Mean CSEM | Cut Score (Scale Score) | CSEM at Cut Score |
| :--- | :--- | :--- | :--- | :--- |
| 9 | 5 | 6.86 | 370 | 6 |
| 10 | 1 | 8.11 |  |  |
| 10 | 2 | 5.35 | 334 | 7 |
| 10 | 3 | 5.20 | 350 | 6 |
| 10 | 4 | 5.81 | 362 | 6 |
| 10 | 5 | 7.43 | 378 | 7 |

Table 22: Achievement Levels and Associated Conditional Standard Error of Measurement (Mathematics)

| Grade | Achievement Level | Mean CSEM | Cut Score (Scale Score) | CSEM at Cut Score |
| :---: | :---: | :---: | :---: | :---: |
| 3 | 1 | 5.80 |  |  |
| 3 | 2 | 4.00 | 285 | 4 |
| 3 | 3 | 4.33 | 297 | 4 |
| 3 | 4 | 5.47 | 311 | 5 |
| 3 | 5 | 9.33 | 327 | 7 |
| 4 | 1 | 6.83 |  |  |
| 4 | 2 | 4.26 | 299 | 5 |
| 4 | 3 | 4.11 | 310 | 4 |
| 4 | 4 | 5.33 | 325 | 5 |
| 4 | 5 | 10.01 | 340 | 7 |
| 5 | 1 | 7.82 |  |  |
| 5 | 2 | 4.33 | 306 | 5 |
| 5 | 3 | 4.00 | 320 | 4 |
| 5 | 4 | 4.58 | 334 | 4 |
| 5 | 5 | 7.86 | 350 | 5 |
| 6 | 1 | 8.43 |  |  |
| 6 | 2 | 4.94 | 310 | 5 |
| 6 | 3 | 4.23 | 325 | 5 |
| 6 | 4 | 4.14 | 339 | 4 |
| 6 | 5 | 6.75 | 356 | 5 |
| 7 | 1 | 10.54 |  |  |
| 7 | 2 | 4.85 | 316 | 6 |
| 7 | 3 | 4.00 | 330 | 4 |
| 7 | 4 | 4.15 | 346 | 4 |


| Grade | Achievement <br> Level | Mean CSEM | Cut Score (Scale Score) | CSEM at Cut Score |
| :--- | :--- | :--- | :--- | :--- |
| 7 | 5 | 5.69 | 360 | 5 |
| 8 | 1 | 9.82 |  |  |
| 8 | 2 | 6.26 | 322 | 7 |
| 8 | 3 | 5.09 | 337 | 6 |
| 8 | 4 | 4.53 | 353 | 5 |
| 8 | 5 | 5.34 | 365 | 5 |

Table 23: Achievement Levels and Associated Conditional Standard Error of Measurement (EOC)

| Grade | Achievement <br> Level | Mean CSEM | Cut Score (Scale Score) | CSEM at Cut Score |
| :--- | :--- | :--- | :--- | :--- |
| Algebra_1 | 1 | 16.06 |  |  |
| Algebra_1 | 2 | 8.12 | 487 | 9 |
| Algebra_1 | 3 | 6.30 | 497 | 7 |
| Algebra_1 | 4 | 5.04 | 518 | 5 |
| Algebra_1 | 5 | 5.05 | 532 | 5 |
| Geometry | 1 | 16.02 |  | 9 |
| Geometry | 2 | 7.81 | 486 | 4 |
| Geometry | 3 | 5.39 | 499 | 4 |
| Geometry | 4 | 4.00 | 521 | 7 |
| Geometry | 5 | 4.44 |  |  |

### 3.6 Writing Prompts Inter-Rater Reliability

Two human raters handscored all writing prompts in grade 10 , and $15 \%$ of grades 4 through 7 prompts. For the online tests, prompts were scored by one human rater, and American Institutes for Research's (AIR) scoring engine was used to provide the second score.

The basic method to compute inter-rater reliability is percentage agreement. As seen in Table 24, the percentage of exact agreement (when two raters gave the same score), the percentage of adjacent ratings (when the difference between two raters was 1), and the percentage of nonadjacent ratings (when the difference was larger than 1) were all computed. In this example, the exact agreement was $2 / 4,50 \%$, and the adjacent and non-adjacent percentages were $25 \%$ each.

Table 24: Percentage Agreement Example

| Response | Rater 1 | Rater 2 | Agreement |
| :--- | :--- | :--- | :--- |
| 1 | 2 | 3 | 1 |
| 2 | 1 | 1 | 0 |
| 3 | 2 | 2 | 0 |
| 4 | 2 | 0 | 2 |

Likewise, inter-rater reliability monitors how often scorers are in exact agreement with each other and ensures that an acceptable agreement rate is maintained. The calculations for inter-rater reliability in this report are as follows:

- Percent Exact: total number of responses by scorer in which scores are equal divided by the number of responses that were scored twice
- Percent Adjacent: total number of responses by scorer in which scores are one score point apart divided by the number of responses that were scored twice
- Percent Non-Adjacent: total number of responses by scorer where scores are more than one score point apart divided by the number of responses that were scored twice, when applicable

Table 25 displays rater-agreement percentages. The percentage of exact agreement between two raters ranged from $71 \%$ to $87 \%$. The percentage of adjacent rating was between $12 \%$ and $29 \%$. The non-adjacent percentages fell between $0 \%$ and $1 \%$. The total number of processed responses does not necessarily correspond to the number of students participating in the Writing portion. These numbers could potentially be higher, as some students are scored more than once when rescoring for some responses, as requested.

Table 25: Inter-Rater Reliability

| Grade | Dimension | \% Exact | \% Adjacent | \% Not Adjacent | Total Number of Processed Responses |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 4 | Purpose, Focus, \& Organization | 79 | 21 | 0 | 262,998 |
|  | Evidence \& Elaboration | 81 | 19 | 0 |  |
|  | Conventions | 78 | 22 | 0 |  |
| 5 | Purpose, Focus, \& Organization | 78 | 21 | 0 | 254,206 |
|  | Evidence \& Elaboration | 78 | 21 | 0 |  |
|  | Conventions | 81 | 19 | 0 |  |
| 6 | Purpose, Focus, \& Organization | 75 | 24 | 1 | 257,039 |
|  | Evidence \& Elaboration | 76 | 24 | 1 |  |
|  | Conventions | 82 | 18 | 0 |  |
| 7 | Purpose, Focus, \& Organization | 78 | 22 | 0 | 247,188 |
|  | Evidence \& Elaboration | 77 | 22 | 1 |  |
|  | Conventions | 87 | 12 | 1 |  |
| 8 | Purpose, Focus, \& Organization | 85 | 14 | 1 | 213,969 |
|  | Evidence \& Elaboration | 86 | 14 | 1 |  |
|  | Conventions | 85 | 15 | 0 |  |
| 9 | Purpose, Focus, \& Organization | 85 | 14 | 1 | 217,882 |
|  | Evidence \& Elaboration | 85 | 14 | 1 |  |
|  | Conventions | 84 | 16 | 0 |  |
| 10 | Purpose, Focus, \& Organization | 71 | 29 | 1 | 400,384 |
|  | Evidence \& Elaboration | 71 | 28 | 1 |  |
|  | Conventions | 80 | 20 | 0 |  |

In addition to inter-rater reliability, validity coefficients, percent exact agreement on validity true scores, and human scores were also calculated. Validity true scores for each dimension were determined by scoring directors, and Test Development Center (TDC) content experts approved those scores. Validity coefficients indicate how often scorers are in exact agreement with previously scored selected responses that are inserted into the scoring queue, and they ensure that an acceptable agreement rate is maintained. The calculations are as follows:

- Percent Exact: total number of responses by scorer where scores are equal divided by the total number of responses that were scored
- Percent Adjacent: total number of responses by scorer where scores are one point apart divided by the total number of responses that were scored
- Percent Non-Adjacent: total number of responses by scorer where scores are more than one score point apart divided by the total number of responses that were scored

Table 26 presents final validity coefficients, which were between 80 and 93.
Table 26: Validity Coefficients

| Grade |  <br> Organization | Evidence \& Elaboration | Conventions |
| :--- | :--- | :--- | :--- |
| 4 | 87 | 87 | 81 |
| 5 | 89 | 85 | 80 |
| 6 | 88 | 88 | 89 |
| 7 | 83 | 82 | 93 |
| 8 | 89 | 90 | 90 |
| 9 | 91 | 90 | 93 |
| 10 | 88 | 88 | 89 |

Cohen's kappa (Cohen, 1968) is an index of inter-rater agreement after accounting for the agreement that could be expected due to chance. This statistic can be computed as

$$
K=\frac{P_{o}-P_{c}}{1-P_{c}}
$$

where $P_{o}$ is the proportion of observed agreement, and $P_{c}$ indicates the proportion of agreement by chance. Cohen's kappa treats all disagreement values with equal weights. Weighted kappa coefficients (Cohen, 1968), however, allow unequal weights, which can be used as a measure of validity. Weighted kappa coefficients were calculated using the formula below:

$$
K_{w}=\frac{P_{o}^{\prime}-P_{c}^{\prime}}{1-P_{c}^{\prime}}
$$

where

$$
\begin{aligned}
P_{o}^{\prime} & =\frac{\sum w_{i j} p_{o i j}}{w_{\max }} \\
P_{c}^{\prime} & =\frac{\sum w_{i j} p_{c i j}}{w_{\max }}
\end{aligned}
$$

where $p_{o i j}$ is the proportion of the judgments observed in the $i j$ th cell, $p_{c i j}$ is the proportion in the $i j$ th cell expected by chance, and $w_{i j}$ is the disagreement weight.

Weighted kappa coefficients for grades 4 through 10 operational writing prompts by dimension are presented in Table 27. They ranged from 0.612 to 0.894 .

Table 27: Weighted Kappa Coefficients

| Grade | Scorer | $\mathbf{N}$ | Purpose, Focus, <br> \& Organization |  <br> Elaboration | Conventions |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 4 | Two Human Raters | 44095 | 0.888 | 0.889 | 0.819 |
| 5 | Two Human Raters | 42693 | 0.894 | 0.878 | 0.842 |
| 6 | Two Human Raters | 41188 | 0.859 | 0.853 | 0.838 |
| 7 | Two Human Raters | 40694 | 0.890 | 0.883 | 0.884 |
| $8^{\star}$ | Machine and Human | 204632 | $0.731(197595)$ | $0.718(197401)$ | $0.612(197830)$ |
| $9^{\star}$ | Machine and Human | 204185 | $0.719(192892)$ | $0.705(193178)$ | $0.644(195525)$ |
| 10 | Two Human Raters | 199370 | 0.769 | 0.765 | 0.740 |

* Grades 8 and 9 kappa values are computed after excluding students with condition codes. N size for each dimension is provided in the parentheses. Appendix C presents contingency tables for these two grades by each dimension.

Grades 8,9 , and 10 Writing prompts were administered online. Grade 10 was scored by two human scorers, while students in grades 8 and 9 received one human score and one machine score through AIR's artificial intelligence (AI) scoring engine.

### 3.6.1 Automated Scoring Engine

AIR's essay scoring engine, Autoscore, uses a statistical process to evaluate Writing prompts. Autoscore evaluates papers against the same rubric used by human raters, but a statistical process is used to analyze each paper and assign scores for each of the three dimensions. The engine uses the same process for scoring essays every time a new prompt is submitted, regardless of whether the data is obtained from an operational assessment or an independent field test.

Statistical rubrics are effectively proxy measures. Although they can directly measure some aspects of Writing conventions (e.g., use of passive voice, misspellings, run-on sentences), they do not directly measure argument structure or content relevance. Hence, although statistical rubrics often prove useful for scoring essays and even for providing some diagnostic feedback in Writing, they do not develop a sufficiently specific model of the correct semantic structure to score many propositional items. Furthermore, they cannot provide the explanatory or diagnostic information available from an explicit rubric. For example, the frequency of incorrect spellings may predict whether a response to a factual item is correct-higher-performing students may also have better spelling skills. Spelling may prove useful in predicting the human score, but it is not the actual reason that the human scorer deducts points. Indeed, statistical rubrics are not about explanation or reason but rather about a prediction of how a human would score the response.

AIR's essay-scoring engine uses a statistical rubric with great success, as measured by the rater agreements observed relative to the human-to-human rater agreements. This technology is similar to all essay-scoring systems in the field. Although some systems replace the statistical process with a "neural network" algorithm, that algorithm functions like the statistical model. Not all descriptions of essay-scoring algorithms are as transparent as AIR's, but whenever a training set is used for the machine to "learn a rubric," the same technology is being used.

The engine is designed to employ a "training set," a set of essays scored with maximally valid scores that are used to form the basis of the prediction model. The quality of the human-assigned scores is critical to the identification of a valid model and final performance of the scoring engine. Moreover, an ideal training sample over-represents higher- and lower-scoring papers and is selected according to a scientific sampling design with known probabilities of selection.

The training process of the scoring engine has two phases. The first phase requires oversampled, high- and low-scoring papers, leaving an equally weighted representative sample for the second phase. The first phase is used to identify concepts that are proportionately represented in higherscoring papers. Here, concepts are defined as words and their synonyms, as well as clusters of words used meaningfully in proximity.

The second phase takes a series of measures on each essay in the remaining training set. These measures include latent semantic analysis (LSA) measures based on the concepts identified in the first phase; other semantic measures indicate the coherence of concepts within and across paragraphs and a range of word-use and syntactic measures. The LSA is similar to a data reduction method identifying common concepts within the narrative and reducing the data to a configurable number of LSA dimensions.

For each trait in the rubric, the system estimates an appropriate statistical model where these LSA and other syntactic characteristics described earlier serve as the independent variables, and the final, resolved score serves as the dependent variable in an ordered probit regression. This model, along with its final parameter estimates, is used to generate a predicted or "proxy" score. The probability of scoring in the $p$ th category is compared to a random draw from the uniform distribution, and a final score point of 1 through 4 is determined from this comparison.

In addition to the training set, an independent random sample of responses is drawn for the crossvalidation of the identified scoring rubric. As with the training set, student responses in the crossvalidation study are handscored, and the LSA and other syntactic characteristics of the papers are computed. Subsequently, a second machine score is generated by applying the model coefficients obtained from the ordered probit in the training set. This forms a predicted score for the papers in the cross-validation set for each dimension in the rubric, which can then be used to evaluate the agreement rates between the human and autoscore engine.

When implementing the scoring engine, we expect that the computer-to-human agreement rates to be at least as high as the human-to-human agreement rates obtained from the double-scored process. If the engine yields scores with rater agreement rates that are at least as high as the human rater agreement rates, then the scoring engine can be deployed for operational scoring. If the computer-to-human agreement rates are not at least as high as the human-to-human rates, then adjustments to the scoring engine statistical model are necessary in order to find a scoring model that yields rater agreement rates that match the human-to-human rates.

To train AIR's AI scoring engine, a subset of papers was scientifically selected and scored by two human raters. Score discrepancies were resolved before being sent from DRC to AIR. The subset was split into a training set with 2,602 or 2,999 papers for grades 8 and 9 , respectively, and the remaining records were used for validation. In addition, in the grade 9 Convention dimension,

2,000 for training and 999 for validation were implemented. The total number of LSA dimensions and the sample size for validation are presented in Table 28. Table 28 also shows that the scoring engine produced comparable results with human scores.

Table 28: Percentage Agreement in Handscoring and Scoring Engine

| Grade | Dimension | Handscoring from DRC |  |  | AIR Scoring Engine |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | \% Exact | \% <br> Adjacent | \% Not Adjacent | LSA | \% Exact | \% <br> Adjacent | \% Not Adjacent | N for Validation |
| 8 | Purpose, Focus, \& Organization | 75.10 | 24.13 | 0.76 | 10 | 75.50 | 23.51 | 0.99 | 302 |
|  | Conventions | 79.00 | 20.73 | 0.27 | 200 | 82.12 | 17.55 | 0.33 | 302 |
|  | Evidence \& Elaboration | 76.09 | 23.11 | 0.80 | 40 | 78.48 | 20.53 | 0.99 | 302 |
| 9 | Purpose, Focus, \& Organization | 79.23 | 20.77 | 0.00 | 50 | 80.76 | 17.84 | 1.40 | 499 |
|  | Conventions | 80.46 | 19.54 | 0.00 | 20 | 79.08 | 18.52 | 2.40 | 999 |
|  | Evidence \& Elaboration | 80.70 | 19.30 | 0.00 | 250 | 76.75 | 20.64 | 2.61 | 499 |

## 4. Evidence of Content Validity

This section demonstrates that the knowledge and skills assessed by the FSA were representative of the content standards of the larger knowledge domain. We describe the content standards for FSA and discuss the test development process, mapping FSA tests to the standards. A complete description of the test development process can be found in Volume 2, Test Development.

### 4.1 Content Standards

The FSA was aligned to the Florida Standards, which were approved by the Florida State Board of Education on February 18, 2014, to be the educational standards for all public schools in the state. The Florida Standards are intended to implement higher standards, with the goal of challenging and motivating Florida's students to acquire stronger critical thinking, problem solving, and communications skills. The Language Arts Florida Standards (LAFS) and the Mathematics Florida Standards (MAFS) are available for review at www.flstandards.org.

Table 29, Table 30, and Table 31 present the reporting categories by grade and test, as well as the number of items measuring each category. Table 32 and Table 33 present the number of items by each reporting category for the accommodated forms.

Table 29: Number of Items for Each ELA Reporting Category

| Reporting Category | Grade* |  |  |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| Key Ideas and Details | 15 | 13 | 15 | 13 | 16 | 15 | 12 | 14 |
| Craft and Structure | 16 | 13 | 13 | 16 | 20 | 19 | 20 | 20 |
| Integration of Knowledge <br> and Ideas | 8 | 13 | 12 | 11 | 8 | 9 | 14 | 13 |
| Language and Editing Task | 11 | 11 | 10 | 12 | 8 | 9 | 8 | 7 |
| Text-Based Writing | 0 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |

* Reporting categories and the number of items belonging to each reporting category are identical for both online and accommodated forms.

Table 30: Number of Items for Each Mathematics Reporting Category

| Grade* | Reporting Category | Number of Items |
| :---: | :---: | :---: |
| 3 | Operations, Algebraic Thinking, and Numbers in Base Ten | 26 |
|  | Numbers and Operations - Fractions | 9 |
|  | Measurement, Data, and Geometry | 19 |
| 4 | Operations and Algebraic Thinking | 11 |
|  | Numbers and Operations in Base Ten | 11 |
|  | Numbers and Operations - Fractions | 14 |
|  | Measurement, Data, and Geometry | 18 |
| 5 | Operations, Algebraic Thinking, and Fractions | 21 |
|  | Numbers and Operations in Base Ten | 15 |
|  | Measurement, Data, and Geometry | 18 |
| 6 | Ratio and Proportional Relationships | 8 |
|  | Expressions and Equations | 17 |
|  | Geometry | 8 |
|  | Statistics and Probability | 11 |
|  | The Number System | 12 |
| 7 | Ratio and Proportional Relationships | 14 |
|  | Expressions and Equations | 12 |
|  | Geometry | 13 |
|  | Statistics and Probability | 9 |
|  | The Number System | 8 |
| 8 | Expressions and Equations | 17 |
|  | Functions | 14 |
|  | Geometry | 15 |
|  | Statistics, Probability, and the Number System | 10 |

[^0] accommodated forms except for grade 8 (Table 32).

Table 31: Number of Items for Each EOC Reporting Category

| Course | Reporting Category | Core Form |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 9 | 10 | 11 | 12 | 13 | 14 | 15 |
| Algebra 1 | Algebra and Modeling |  |  |  | 24 | 24 | 24 | 24 |
|  | Functions and Modeling |  |  |  | 23 | 23 | 23 | 23 |
|  | Statistics and the Number System |  |  |  | 11 | 11 | 11 | 11 |
| Geometry | Congruence, Similarity, Right Triangles, and Trigonometry | 27 | 27 | 27 |  |  |  |  |
|  | Circles, Geometric Measurement, and Geometric Properties with Equations | 22 | 22 | 22 |  |  |  |  |
|  | Modeling with Geometry | 9 | 9 | 9 |  |  |  |  |

Table 32: Number of Items for Each Mathematics Accommodated Reporting Category

| Grade | Reporting Category | Number of Items |
| :---: | :--- | :--- |
| 8 | Expressions and Equations | 17 |
|  | Functions | 14 |
|  | Geometry | 16 |
|  | Statistics, Probability, and the Number System | 9 |

Table 33: Number of Items for Each EOC Accommodated Reporting Category

| Course | Reporting Category | Number of <br> Items |
| :--- | :--- | :--- |
|  | Algebra and Modeling | 26 |
|  | Functions and Modeling | 25 |
|  | Statistics and the Number System | 7 |
| Geometry | Congruence, Similarity, Right Triangles, and Trigonometry | 27 |
|  | Circles, Geometric Measurement, and Geometric Properties with Equations | 22 |
|  | Modeling with Geometry | 9 |

### 4.2 TEST SPECIFICATIONS

Blueprints were developed to ensure that the test and the items were aligned to the prioritized standards that they were intended to measure. For more detail, please see Volume 2, Section 2. The Florida Standards Assessments (FSA) were composed of test items that included traditional multiple-choice items, items that required students to type or write a response, and technologyenhanced items (TEI). Technology-enhanced items are computer-delivered items that require students to interact with test content to select, construct, and support their answers. The blueprints
specified the percentage of operational items that were to be administered. The blueprints also included the minimum and maximum number of items for each of the reporting categories, and constraints on selecting items for the depth of knowledge (DoK) levels in Reading. The minimum and maximum number of items by grade and subject and other details on the blueprint are presented in appendices of Volume 2.

### 4.3 Test Development

For the 2018 Florida Standards Assessments administration, the American Institutes for Research in collaboration with the Florida Department of Education and its Test Development Center (TDC), constructed test forms for ELA grades 3 through 10 and grade 10 retake, Mathematics grades 3 through 8, and End-of-Course Assessments (Algebra 1, Geometry).

Test construction began during the summer of 2017, when all parties met face to face to select items that aligned to the FSA standards and blueprints designed for the FSA. Curricular, psychometric, and policy experts constructed test forms carefully, evaluating the fit of each item's statistical characteristics and the alignment of the item to Florida's standards. The content guidelines, which describe standards coverage and item type coverage, are outlined in detail in Appendices A and B of Volume 2, Test Development.

The Florida Standards Assessments item pool grows each year by field testing new items. Any item used on an assessment was field tested before it was used as an operational item. Field testing was conducted during the spring as part of the regular administration. The field-test items utilized the same positions as anchor items. In order to keep the test length consistent, placeholder items were placed into the field-test positions on some of the forms. The number of forms constructed for a given grade and subject was at most 40, including field-test and anchor forms.

After operational forms were developed, the AIR and TDC content specialists worked together to assign newly developed items to field test forms for field testing. The teams addressed the following factors when embedding field-test items into operational test forms for the spring administration:

- Ensured field-test items did not cue or clue answers to other field-test items on the form
- Ensured field-test items that cued or clued answers to operational items were not field tested
- Included a mix of items covering multiple reporting categories and standards on each form
- Selected items in the field-test sets that reflected a range of difficulty levels and cognitive levels
- Minimized abrupt transitions from one subject strand or mental construct to another
- Selected items that were needed for appropriate standard coverage in the item bank-
- Selected items that were needed for appropriate format variety in the item bank
- Maintained awareness of the distribution of keys and the number of adjacent items having the same key


### 4.4 ALIGNMENT OF FSA ITEM BANKS TO THE CONTENT STANDARDS AND BENCHMARKS

A third-party, independent alignment study was completed. The study found that items were fully aligned with the intended content and that items in FSA test forms demonstrated a good representation of the standards-the Language Arts Florida Standards (LAFS) and the Mathematics Florida Standards (MAFS). A full report on alignment was provided in Volume 4 Appendix D of 2015-2016 FSA Annual Technical Report.

## 5. Evidence on Internal-External Structure

In this section, we explore the internal structure of the assessment using the scores provided at the reporting category level. The relationship of the subscores is just one indicator of the test dimensionality.

In ELA grades 4 through 10, there are five reporting categories per grade: Key Ideas and Details, Craft and Structure, Integration of Knowledge and Ideas, Language and Editing Task, and TextBased Writing. Reading grade 3 has the same reporting categories, with the exception of TextBased Writing. In Mathematics and EOC tests, reporting categories differ in each grade or course (see Table 29, Table 30, and Table 31 for reporting category information).

Raw scores based on each reporting category were provided to students. Evidence is needed to verify that the raw score for each reporting category provides both different and useful information for student achievement.

It may not be reasonable to expect that the reporting category scores are completely orthogonalthis would suggest that there are no relationships among reporting category scores and would make justification of a unidimensional IRT model difficult, although we could then easily justify reporting these separate scores. On the contrary, if the reporting categories were perfectly correlated, we could justify a unidimensional model, but we could not justify the reporting of separate scores.

One pathway to explore the internal structure of the test is via a second-order factor model, assuming a general mathematics construct (first factor) with reporting categories (second factor), and that the items load onto the reporting category they intend to measure. If the first-order factors are highly correlated and the model fits data well for the second-order model, this provides evidence of unidimensionality as well as reporting subscores.

Another pathway is to explore observed correlations between the subscores. However, as each reporting category is measured with a small number of items, the standard errors of the observed scores within each reporting category are typically larger than the standard error of the total test score. Disattenuating for measurement error could offer some insight into the theoretical true score correlations. Both observed correlations and disattenuated correlations are provided in the following section.

### 5.1 Correlations Among Reporting Category Scores

Table 34 through Table 36 present the observed correlation matrix of the reporting category raw scores for each subject area. In ELA, the correlations among the reporting categories range from 0.43 to 0.78 . The Language and Editing Task items and Text-Based Writing items exhibited slightly lower correlations, with the other reporting categories ranging from 0.45 to 0.63 . For Mathematics, the correlations were between 0.65 and 0.85 . For EOC, the correlations between the three subscales fell between 0.69 and 0.86 . Observed correlations from the accommodated forms are presented in Table 37 through Table 39. The correlations varied between 0.41 and 0.77 for ELA, 0.55 and 0.84 for Mathematics, and 0.56 and 0.74 for EOC.

In some instances, these correlations were lower than one might expect. However, as previously noted, the correlations were subject to a large amount of measurement error at the strand level,
given the limited number of items from which the scores were derived. Consequently, overinterpretation of these correlations, as either high or low, should be made cautiously, which the Florida Department of Education cautions each year when scores are released.

Table 40 through Table 45 display disattenuated correlations. Disattenuated values greater than 1.00 are reported as $1.00^{*}$. In ELA, the Writing dimension had the lowest correlations among the five reporting categories. For the Writing dimension, the average value was 0.71 , and the minimum was 0.64 , whereas the overall average disattenuated correlation for ELA was 0.82 .

Table 34: Observed Correlation Matrix Among Reporting Categories (ELA)

| Grade | Reporting Category | Number of Items | Cat1 | Cat2 | Cat3 | Cat4 | Cat5 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 3 | Key Ideas and Details (Cat1) | 15 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 16 | 0.77 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 8 | 0.65 | 0.66 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 11 | 0.61 | 0.61 | 0.51 | 1.00 |  |
| 4 | Key Ideas and Details (Cat1) | 13 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 13 | 0.67 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 13 | 0.68 | 0.61 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 11 | 0.60 | 0.57 | 0.55 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 1 | 0.51 | 0.49 | 0.46 | 0.51 | 1.00 |
| 5 | Key Ideas and Details (Cat1) | 15 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 13 | 0.73 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 12 | 0.78 | 0.72 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 10 | 0.59 | 0.57 | 0.58 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 1 | 0.55 | 0.53 | 0.54 | 0.51 | 1.00 |
| 6 | Key Ideas and Details (Cat1) | 13 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 16 | 0.74 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 11 | 0.71 | 0.67 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 12 | 0.61 | 0.61 | 0.55 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 1 | 0.52 | 0.51 | 0.49 | 0.54 | 1.00 |
| 7 | Key Ideas and Details (Cat1) | 16 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 20 | 0.75 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 8 | 0.65 | 0.65 | 1.00 |  |  |


| Language and Editing Task <br> (Cat4) | 8 | 0.51 | 0.51 | 0.43 | 1.00 |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | Text-Based Writing (Cat5) | 1 | 0.58 | 0.57 | 0.49 | 0.45 | 1.00 |
|  | Key Ideas and Details (Cat1) | 15 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 19 | 0.75 | 1.00 |  |  |  |
|  | Integration of Knowledge and <br> Ideas (Cat3) | 9 | 0.68 | 0.72 | 1.00 |  |  |
| Language and Editing Task <br> (Cat4) | 9 | 0.54 | 0.61 | 0.52 | 1.00 |  |  |
|  | Text-Based Writing (Cat5) | 1 | 0.56 | 0.63 | 0.55 | 0.55 | 1.00 |
|  | Key Ideas and Details (Cat1) | 12 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 20 | 0.76 | 1.00 |  |  |  |
| Integration of Knowledge and <br> Ideas (Cat3) | 14 | 0.74 | 0.74 | 1.00 |  |  |  |
| Language and Editing Task <br> (Cat4) | 8 | 0.55 | 0.56 | 0.54 | 1.00 |  |  |
|  | Text-Based Writing (Cat5) | 1 | 0.60 | 0.62 | 0.58 | 0.51 | 1.00 |
|  | Key Ideas and Details (Cat1) | 14 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 20 | 0.75 | 1.00 |  |  |  |
| Integration of Knowledge and <br> Ideas (Cat3) | 13 | 0.72 | 0.71 | 1.00 |  |  |  |
| Language and Editing Task <br> (Cat4) | 7 | 0.57 | 0.57 | 0.55 | 1.00 | 1.00 |  |
| Text-Based Writing (Cat5) | 1 | 0.62 | 0.62 | 0.59 | 0.54 |  |  |

Table 35: Observed Correlation Matrix Among Reporting Categories (Mathematics)

| Grade | Reporting Category | Number <br> of Items | Cat1 | Cat2 | Cat3 | Cat4 | Cat5 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | Operations, Algebraic Thinking, <br> and Numbers in Base Ten <br> (Cat1) | 26 | 1.00 |  |  |  |  |
|  | Numbers and Operations - <br> Fractions (Cat2) | 9 | 0.72 | 1.00 |  |  |  |
|  | Measurement, Data, and <br> Geometry (Cat3) | 19 | 0.83 | 0.75 | 1.00 |  |  |
| 4 | Operations and Algebraic <br> Thinking (Cat1) | 11 | 1.00 |  |  |  |  |
|  | 11 | 0.74 | 1.00 |  |  |  |  |
|  | 14 | 0.77 | 0.79 | 1.00 | 1.00 | 0.83 | 0.78 |
| Measurement, Data, and <br> Geometry (Cat4) | 18 | 0.77 | 0.75 |  |  |  |  |


| 5 | Operations, Algebraic Thinking, and Fractions (Cat1) | 21 | 1.00 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Numbers and Operations in Base Ten (Cat2) | 15 | 0.85 | 1.00 |  |  |  |
|  | Measurement, Data, and Geometry (Cat3) | 18 | 0.84 | 0.81 | 1.00 |  |  |
| 6 | Ratio and Proportional Relationships (Cat1) | 8 | 1.00 |  |  |  |  |
|  | Expressions and Equations (Cat2) | 17 | 0.77 | 1.00 |  |  |  |
|  | Geometry (Cat3) | 8 | 0.68 | 0.75 | 1.00 |  |  |
|  | Statistics and Probability (Cat4) | 11 | 0.70 | 0.75 | 0.69 | 1.00 |  |
|  | The Number System (Cat5) | 12 | 0.76 | 0.84 | 0.74 | 0.73 | 1.00 |
| 7 | Ratio and Proportional Relationships (Cat1) | 14 | 1.00 |  |  |  |  |
|  | Expressions and Equations (Cat2) | 12 | 0.76 | 1.00 |  |  |  |
|  | Geometry (Cat3) | 13 | 0.76 | 0.73 | 1.00 |  |  |
|  | Statistics and Probability (Cat4) | 9 | 0.75 | 0.67 | 0.69 | 1.00 |  |
|  | The Number System (Cat5) | 8 | 0.77 | 0.74 | 0.74 | 0.69 | 1.00 |
| 8 | Expressions and Equations (Cat1) | 17 | 1.00 |  |  |  |  |
|  | Functions (Cat2) | 14 | 0.70 | 1.00 |  |  |  |
|  | Geometry (Cat3) | 15 | 0.70 | 0.66 | 1.00 |  |  |
|  | Statistics and Probability and the Number System (Cat4) | 10 | 0.72 | 0.65 | 0.65 | 1.00 |  |

Table 36: Observed Correlation Matrix Among Reporting Categories (EOC)

| Course/Form | Reporting Category | Number of Items | Cat1 | Cat2 | Cat3 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Algebra 1/Core 12 | Algebra and Modeling (Cat1) | 24 | 1.00 |  |  |
|  | Functions and Modeling (Cat2) | 23 | 0.82 | 1.00 |  |
|  | Statistics and the Number System (Cat3) | 11 | 0.76 | 0.74 | 1.00 |
| Algebra 1/Core 13 | Algebra and Modeling (Cat1) | 24 | 1.00 |  |  |
|  | Functions and Modeling (Cat2) | 23 | 0.83 | 1.00 |  |
|  | Statistics and the Number System (Cat3) | 11 | 0.74 | 0.73 | 1.00 |
| Algebra 1/Core 14 | Algebra and Modeling (Cat1) | 24 | 1.00 |  |  |
|  | Functions and Modeling (Cat2) | 23 | 0.83 | 1.00 |  |
|  | Statistics and the Number System (Cat3) | 11 | 0.72 | 0.70 | 1.00 |


| Algebra 1/Core 15 | Algebra and Modeling (Cat1) | 24 | 1.00 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Functions and Modeling (Cat2) | 23 | 0.82 | 1.00 |  |
|  | Statistics and the Number System (Cat3) | 11 | 0.71 | 0.69 | 1.00 |
| Geometry/Core 9 | Congruence, Similarity, Right Triangles, and Trigonometry (Cat1) | 27 | 1.00 |  |  |
|  | Circles, Geometric Measurement, and Geometric Properties with Equations (Cat2) | 22 | 0.83 | 1.00 |  |
|  | Modeling with Geometry (Cat3) | 9 | 0.77 | 0.78 | 1.00 |
| Geometry/Core 10 | Congruence, Similarity, Right Triangles, and Trigonometry (Cat1) | 27 | 1.00 |  |  |
|  | Circles, Geometric Measurement, and Geometric Properties with Equations (Cat2) | 22 | 0.85 | 1.00 |  |
|  | Modeling with Geometry (Cat3) | 9 | 0.80 | 0.81 | 1.00 |
| Geometry/Core 11 | Congruence, Similarity, Right Triangles, and Trigonometry (Cat1) | 27 | 1.00 |  |  |
|  | Circles, Geometric Measurement, and Geometric Properties with Equations (Cat2) | 22 | 0.86 | 1.00 |  |
|  | Modeling with Geometry (Cat3) | 9 | 0.79 | 0.79 | 1.00 |

Table 37: Observed Correlation Matrix Among Reporting Categories (ELA

## Accommodated Forms)

| Grade | Reporting Category | Number <br> of Items | Cat1 | Cat2 | Cat3 | Cat4 | Cat5 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | Key Ideas and Details (Cat1) | 13 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 13 | 0.63 | 1.00 |  |  |  |
|  | Integration of Knowledge and <br> Ideas (Cat3) | 13 | 0.62 | 0.57 | 1.00 |  |  |
|  | Language and Editing Task <br> (Cat4) | 11 | 0.53 | 0.55 | 0.50 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 1 | 0.48 | 0.42 | 0.41 | 0.49 | 1.00 |
| 5 | Key Ideas and Details (Cat1) | 15 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 13 | 0.69 | 1.00 |  |  |  |
|  | 12 | 0.75 | 0.67 | 1.00 |  |  |  |
|  | Language and Editing Task <br> (Cat4) | 10 | 0.53 | 0.52 | 0.51 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 1 | 0.50 | 0.45 | 0.50 | 0.47 | 1.00 |
| 6 | Key Ideas and Details (Cat1) | 13 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 16 | 0.71 | 1.00 |  |  |  |


|  | Integration of Knowledge and Ideas (Cat3) | 11 | 0.64 | 0.61 | 1.00 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Language and Editing Task (Cat4) | 12 | 0.53 | 0.51 | 0.45 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 1 | 0.53 | 0.49 | 0.47 | 0.51 | 1.00 |
| 7 | Key Ideas and Details (Cat1) | 16 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 20 | 0.72 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 8 | 0.61 | 0.63 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 8 | 0.52 | 0.52 | 0.44 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 1 | 0.56 | 0.56 | 0.46 | 0.43 | 1.00 |
| 8 | Key Ideas and Details (Cat1) | 15 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 19 | 0.74 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 9 | 0.67 | 0.71 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 9 | 0.54 | 0.58 | 0.48 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 1 | 0.49 | 0.56 | 0.47 | 0.49 | 1.00 |
| 9 | Key Ideas and Details (Cat1) | 12 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 20 | 0.76 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 14 | 0.72 | 0.72 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 8 | 0.55 | 0.56 | 0.53 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 1 | 0.53 | 0.55 | 0.51 | 0.52 | 1.00 |
| 10 | Key Ideas and Details (Cat1) | 14 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 20 | 0.77 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 13 | 0.72 | 0.71 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 7 | 0.59 | 0.58 | 0.53 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 1 | 0.57 | 0.60 | 0.55 | 0.51 | 1.00 |

Table 38: Observed Correlation Matrix Among Reporting Categories (Mathematics Accommodated Forms)

| Grade | Reporting Category | Number <br> of Items | Cat1 | Cat2 | Cat3 | Cat4 | Cat5 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 3 | Operations, Algebraic Thinking, <br> and Numbers in Base Ten <br> (Cat1) | 26 | 1.00 |  |  |  |  |
|  | Numbers and Operations - <br> Fractions (Cat2) | 9 | 0.69 | 1.00 |  |  |  |


|  | Measurement, Data, and Geometry (Cat3) | 19 | 0.84 | 0.72 | 1.00 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 4 | Operations and Algebraic Thinking (Cat1) | 11 | 1.00 |  |  |  |  |
|  | Numbers and Operations in Base Ten (Cat2) | 11 | 0.77 | 1.00 |  |  |  |
|  | Numbers and Operations Fractions (Cat3) | 14 | 0.78 | 0.80 | 1.00 |  |  |
|  | Measurement, Data, and Geometry (Cat4) | 18 | 0.80 | 0.80 | 0.83 | 1.00 |  |
| 5 | Operations, Algebraic Thinking, and Fractions (Cat1) | 21 | 1.00 |  |  |  |  |
|  | Numbers and Operations in Base Ten (Cat2) | 15 | 0.84 | 1.00 |  |  |  |
|  | Measurement, Data, and Geometry (Cat3) | 18 | 0.80 | 0.80 | 1.00 |  |  |
| 6 | Ratio and Proportional Relationships (Cat1) | 8 | 1.00 |  |  |  |  |
|  | Expressions and Equations (Cat2) | 17 | 0.73 | 1.00 |  |  |  |
|  | Geometry (Cat3) | 8 | 0.63 | 0.64 | 1.00 |  |  |
|  | Statistics and Probability (Cat4) | 11 | 0.66 | 0.66 | 0.58 | 1.00 |  |
|  | The Number System (Cat5) | 12 | 0.75 | 0.78 | 0.61 | 0.67 | 1.00 |
| 7 | Ratio and Proportional Relationships (Cat1) | 14 | 1.00 |  |  |  |  |
|  | Expressions and Equations (Cat2) | 12 | 0.72 | 1.00 |  |  |  |
|  | Geometry (Cat3) | 13 | 0.73 | 0.69 | 1.00 |  |  |
|  | Statistics and Probability (Cat4) | 9 | 0.72 | 0.64 | 0.67 | 1.00 |  |
|  | The Number System (Cat5) | 8 | 0.72 | 0.70 | 0.72 | 0.68 | 1.00 |
| 8 | Expressions and Equations (Cat1) | 17 | 1.00 |  |  |  |  |
|  | Functions (Cat2) | 14 | 0.70 | 1.00 |  |  |  |
|  | Geometry (Cat3) | 16 | 0.71 | 0.67 | 1.00 |  |  |
|  | Statistics and Probability and the Number System (Cat4) | 9 | 0.60 | 0.59 | 0.55 | 1.00 |  |

Table 39: Observed Correlation Matrix Among Reporting Categories (EOC Accommodated Forms)

| Course | Reporting Category | Number of <br> Items | Cat1 | Cat2 | Cat3 |
| :--- | :--- | :--- | :--- | :--- | :---: |
| Algebra 1 | Algebra and Modeling (Cat1) | 26 | 1.00 |  |  |
|  | Functions and Modeling (Cat2) | 25 | 0.74 | 1.00 |  |


|  | Statistics and the Number System <br> (Cat3) | 7 | 0.60 | 0.56 | 1.00 |
| :--- | :--- | :--- | :--- | :--- | :--- |
| Geometry | Congruence, Similarity, Right <br> Triangles, and Trigonometry (Cat1) | 27 | 1.00 |  |  |
|  | 22 | 0.70 | 1.00 |  |  |
|  | 9 | 0.67 | 0.67 | 1.00 |  |

Table 40: Disattenuated Correlation Matrix Among Reporting Categories (ELA)

| Grade | Reporting Category | Number of Items | Cat1 | Cat2 | Cat3 | Cat4 | Cat5 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 3 | Key Ideas and Details (Cat1) | 15 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 16 | 1.00 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 8 | 0.97 | 0.98 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 11 | 0.85 | 0.86 | 0.83 | 1.00 |  |
| 4 | Key Ideas and Details (Cat1) | 13 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 13 | 0.95 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 13 | 0.99 | 0.96 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 11 | 0.87 | 0.90 | 0.87 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 1 | 0.65 | 0.67 | 0.64 | 0.71 | 1.00 |
| 5 | Key Ideas and Details (Cat1) | 15 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 13 | 0.98 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 12 | 1.00* | 1.00 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 10 | 0.88 | 0.90 | 0.89 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 1 | 0.68 | 0.70 | 0.69 | 0.73 | 1.00 |
| 6 | Key Ideas and Details (Cat1) | 13 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 16 | 1.00 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 11 | 1.00 | 0.97 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 12 | 0.86 | 0.89 | 0.83 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 1 | 0.66 | 0.66 | 0.66 | 0.73 | 1.00 |
| 7 | Key Ideas and Details (Cat1) | 16 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 20 | 0.99 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 8 | 0.99 | 0.99 | 1.00 |  |  |


|  | Language and Editing Task (Cat4) | 8 | 0.86 | 0.85 | 0.83 | 1.00 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Text-Based Writing (Cat5) | 1 | 0.73 | 0.71 | 0.71 | 0.71 | 1.00 |
| 8 | Key Ideas and Details (Cat1) | 15 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 19 | 1.00 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 9 | 0.99 | 0.98 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 9 | 0.86 | 0.91 | 0.86 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 1 | 0.72 | 0.76 | 0.73 | 0.80 | 1.00 |
| 9 | Key Ideas and Details (Cat1) | 12 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 20 | 1.00* | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 14 | 1.00* | 0.99 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 8 | 0.80 | 0.80 | 0.80 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 1 | 0.75 | 0.76 | 0.74 | 0.68 | 1.00 |
| 10 | Key Ideas and Details (Cat1) | 14 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 20 | 1.00* | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 13 | 0.97 | 0.97 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 7 | 0.87 | 0.87 | 0.85 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 1 | 0.77 | 0.77 | 0.73 | 0.76 | 1.00 |

Table 41: Disattenuated Correlation Matrix Among Reporting Categories (Mathematics)

| Grade | Reporting Category | Number <br> of Items | Cat1 | Cat2 | Cat3 | Cat4 | Cat5 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | Operations, Algebraic Thinking, <br> and Numbers in Base Ten <br> (Cat1) | 26 | 1.00 |  |  |  |  |
| Numbers and Operations - <br> Fractions (Cat2) | 9 | 0.90 | 1.00 |  |  |  |  |
| Measurement, Data, and <br> Geometry (Cat3) | 19 | 0.95 | 0.96 | 1.00 |  |  |  |
| 4 | Operations and Algebraic <br> Thinking (Cat1) | 11 | 1.00 |  |  |  |  |
|  | 11 | 0.98 | 1.00 |  |  |  |  |
|  | 14 | 0.99 | 0.98 | 1.00 | 1.00 |  |  |
| Measurement, Data, and <br> Geometry (Cat4) | 18 | 0.97 | 0.95 | 0.98 |  |  |  |


| 5 | Operations, Algebraic Thinking, and Fractions (Cat1) | 21 | 1.00 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Numbers and Operations in Base Ten (Cat2) | 15 | 0.97 | 1.00 |  |  |  |
|  | Measurement, Data, and Geometry (Cat3) | 18 | 0.97 | 0.94 | 1.00 |  |  |
| 6 | Ratio and Proportional Relationships (Cat1) | 8 | 1.00 |  |  |  |  |
|  | Expressions and Equations (Cat2) | 17 | 0.99 | 1.00 |  |  |  |
|  | Geometry (Cat3) | 8 | 0.94 | 0.94 | 1.00 |  |  |
|  | Statistics and Probability (Cat4) | 11 | 0.96 | 0.95 | 0.92 | 1.00 |  |
|  | The Number System (Cat5) | 12 | 1.00* | 1.00* | 0.95 | 0.94 | 1.00 |
| 7 | Ratio and Proportional Relationships (Cat1) | 14 | 1.00 |  |  |  |  |
|  | Expressions and Equations (Cat2) | 12 | 0.96 | 1.00 |  |  |  |
|  | Geometry (Cat3) | 13 | 0.92 | 0.95 | 1.00 |  |  |
|  | Statistics and Probability (Cat4) | 9 | 0.99 | 0.96 | 0.93 | 1.00 |  |
|  | The Number System (Cat5) | 8 | 0.98 | 1.00* | 0.96 | 0.99 | 1.00 |
| 8 | Expressions and Equations (Cat1) | 17 | 1.00 |  |  |  |  |
|  | Functions (Cat2) | 14 | 0.94 | 1.00 |  |  |  |
|  | Geometry (Cat3) | 15 | 0.92 | 0.93 | 1.00 |  |  |
|  | Statistics and Probability and the Number System (Cat4) | 10 | 0.96 | 0.93 | 0.91 | 1.00 |  |

Table 42: Disattenuated Correlation Matrix Among Reporting Categories (EOC)

| Course/Form | Reporting Category | Number of Items | Cat1 | Cat2 | Cat3 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Algebra 1/Core 12 | Algebra and Modeling (Cat1) | 24 | 1.00 |  |  |
|  | Functions and Modeling (Cat2) | 23 | 0.98 | 1.00 |  |
|  | Statistics and the Number System (Cat3) | 11 | 0.99 | 1.00* | 1.00 |
| Algebra 1/Core 13 | Algebra and Modeling (Cat1) | 24 | 1.00 |  |  |
|  | Functions and Modeling (Cat2) | 23 | 0.98 | 1.00 |  |
|  | Statistics and the Number System (Cat3) | 11 | 0.98 | 1.00 | 1.00 |
| Algebra 1/Core 14 | Algebra and Modeling (Cat1) | 24 | 1.00 |  |  |
|  | Functions and Modeling (Cat2) | 23 | 0.98 | 1.00 |  |
|  | Statistics and the Number System (Cat3) | 11 | 0.97 | 0.99 | 1.00 |


| Algebra 1/Core 15 | Algebra and Modeling (Cat1) | 24 | 1.00 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Functions and Modeling (Cat2) | 23 | 0.98 | 1.00 |  |
|  | Statistics and the Number System (Cat3) | 11 | 0.97 | 0.99 | 1.00 |
| Geometry/Core 6 | Congruence, Similarity, Right Triangles, and Trigonometry (Cat1) | 27 | 1.00 |  |  |
|  | Circles, Geometric Measurement, and Geometric Properties with Equations (Cat2) | 22 | 0.96 | 1.00 |  |
|  | Modeling with Geometry (Cat3) | 9 | 0.97 | 0.96 | 1.00 |
| Geometry/Core 7 | Congruence, Similarity, Right Triangles, and Trigonometry (Cat1) | 27 | 1.00 |  |  |
|  | Circles, Geometric Measurement, and Geometric Properties with Equations (Cat2) | 22 | 0.96 | 1.00 |  |
|  | Modeling with Geometry (Cat3) | 9 | 0.96 | 0.97 | 1.00 |
| Geometry/Core 8 | Congruence, Similarity, Right Triangles, and Trigonometry (Cat1) | 27 | 1.00 |  |  |
|  | Circles, Geometric Measurement, and Geometric Properties with Equations (Cat2) | 22 | 0.97 | 1.00 |  |
|  | Modeling with Geometry (Cat3) | 9 | 0.97 | 0.96 | 1.00 |

Table 43: Disattenuated Correlation Matrix Among Reporting Categories (ELA Accommodated Forms)

| Grade | Reporting Category | Number <br> of Items | Cat1 | Cat2 | Cat3 | Cat4 | Cat5 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | Key Ideas and Details (Cat1) | 13 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 13 | 0.95 | 1.00 |  |  |  |
|  | Integration of Knowledge and <br> Ideas (Cat3) | 13 | 0.98 | 0.98 | 1.00 |  |  |
|  | Language and Editing Task <br> (Cat4) | 11 | 0.77 | 0.86 | 0.81 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 1 | 0.62 | 0.59 | 0.60 | 0.66 | 1.00 |
| 5 | Key Ideas and Details (Cat1) | 15 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 13 | 0.97 | 1.00 |  |  |  |
|  | Integration of Knowledge and <br> Ideas (Cat3) | 12 | $1.00^{*}$ | 0.99 | 1.00 |  |  |
| Language and Editing Task <br> (Cat4) | 10 | 0.80 | 0.83 | 0.79 | 1.00 |  |  |


|  | Text-Based Writing (Cat5) | 1 | 0.62 | 0.60 | 0.65 | 0.66 | 1.00 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | Key Ideas and Details (Cat1) | 13 | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 16 | $1.00^{*}$ | 1.00 |  |  |  |


|  | Language and Editing Task <br> (Cat4) | 7 | 0.89 | 0.89 | 0.83 | 1.00 |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | Text-Based Writing (Cat5) | 1 | 0.70 | 0.75 | 0.70 | 0.71 | 1.00 |

Table 44: Disattenuated Correlation Matrix Among Reporting Categories (Mathematics Accommodated Forms)

| Grade | Reporting Category | Number of Items | Cat1 | Cat2 | Cat3 | Cat4 | Cat5 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 3 | Operations, Algebraic Thinking, and Numbers in Base Ten (Cat1) | 26 | 1.00 |  |  |  |  |
|  | Numbers and Operations Fractions (Cat2) | 9 | 0.89 | 1.00 |  |  |  |
|  | Measurement, Data, and Geometry (Cat3) | 19 | 0.95 | 0.96 | 1.00 |  |  |
| 4 | Operations and Algebraic Thinking (Cat1) | 11 | 1.00 |  |  |  |  |
|  | Numbers and Operations in Base Ten (Cat2) | 11 | 0.98 | 1.00 |  |  |  |
|  | Numbers and Operations Fractions (Cat3) | 14 | 0.97 | 0.97 | 1.00 |  |  |
|  | Measurement, Data, and Geometry (Cat4) | 18 | 0.99 | 0.96 | 0.97 | 1.00 |  |
| 5 | Operations, Algebraic Thinking, and Fractions (Cat1) | 21 | 1.00 |  |  |  |  |
|  | Numbers and Operations in Base Ten (Cat2) | 15 | 0.97 | 1.00 |  |  |  |
|  | Measurement, Data, and Geometry (Cat3) | 18 | 0.97 | 0.95 | 1.00 |  |  |
| 6 | Ratio and Proportional Relationships (Cat1) | 8 | 1.00 |  |  |  |  |
|  | Expressions and Equations (Cat2) | 17 | 0.96 | 1.00 |  |  |  |
|  | Geometry (Cat3) | 8 | 0.86 | 0.84 | 1.00 |  |  |
|  | Statistics and Probability (Cat4) | 11 | 0.93 | 0.91 | 0.82 | 1.00 |  |
|  | The Number System (Cat5) | 12 | 0.97 | 0.99 | 0.79 | 0.91 | 1.00 |
| 7 | Ratio and Proportional Relationships (Cat1) | 14 | 1.00 |  |  |  |  |
|  | Expressions and Equations (Cat2) | 12 | 0.98 | 1.00 |  |  |  |
|  | Geometry (Cat3) | 13 | 0.94 | 0.96 | 1.00 |  |  |
|  | Statistics and Probability (Cat4) | 9 | 1.00 | 0.95 | 0.95 | 1.00 |  |
|  | The Number System (Cat5) | 8 | 0.94 | 0.98 | 0.96 | 0.98 | 1.00 |


| 8 | Expressions and Equations <br> (Cat1) | 17 | 1.00 |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | Functions (Cat2) | 14 | 0.93 | 1.00 |  |  |  |
|  | Geometry (Cat3) | 16 | 0.93 | 0.94 | 1.00 |  |  |
|  | Statistics and Probability and <br> the Number System (Cat4) | 9 | 0.92 | 0.96 | 0.89 | 1.00 |  |

Table 45: Disattenuated Correlation Matrix Among Reporting Categories (EOC Accommodated Forms)

| Course | Reporting Category | Number of <br> Items | Cat1 | Cat2 | Cat3 |
| :--- | :--- | :--- | :--- | :--- | :---: |
| Algebra 1 | Algebra and Modeling (Cat1) | 26 | 1.00 |  |  |
|  | Functions and Modeling (Cat2) | 25 | $1.00^{*}$ | 1.00 |  |
|  | Statistics and the Number System <br> (Cat3) | 7 | $1.00^{*}$ | $1.00^{*}$ | 1.00 |
| Geometry | Congruence, Similarity, Right <br> Triangles and Trigonometry (Cat1) | 27 | 1.00 |  |  |
|  | 22 | 0.93 | 1.00 | 1.00 |  |
|  | Modeling with Geometry (Cat3) | 9 | 0.96 | 0.95 |  |

### 5.2 Confirmatory Factor Analysis

The FSA had test items designed to measure different standards and higher-level reporting categories. Test scores were reported as an overall performance measure. Additionally, scores on the various reporting categories were also provided as indices of strand-specific performance. The strand scores were reported in a fashion that aligned with the theoretical structure of the test derived from the test blueprint.

The results in this section are intended to provide evidence that the methods for reporting FSA strand scores align with the underlying structure of the test and also provide evidence for appropriateness of the selected IRT models. This section is based on a second-order confirmatory factor analysis, in which the first order factors load onto a common underlying factor. The firstorder factors represent the dimensions of the test blueprint, and items load onto factors they are intended to measure. The underlying structure of the ELA and Mathematics tests was generally common across all grades, which is useful for comparing the results of our analyses across the grades.

While the test consisted of items targeting different standards, all items within a grade and subject were calibrated concurrently using the various IRT models described in this technical report. This implies the pivotal IRT assumption of local independence (Lord, 1980). Formally stated, this assumption posits that the probability of the outcome on item $i$ depends only on the student's ability and the characteristics of the item. Beyond that, the score of item $i$ is independent of the outcome of all other items. From this assumption, the joint density (i.e., the likelihood) is viewed
as the product of the individual densities. Thus, maximum likelihood estimation of person and item parameters in traditional item response theory (IRT) is derived on the basis of this theory.

The measurement model and the score reporting method assume a single underlying factor, with separate factors representing each of the reporting categories. Consequently, it is important to collect validity evidence on the internal structure of the assessment to determine the rationality of conducting concurrent calibrations, as well as using these scoring and reporting methods.

### 5.2.1 Factor Analytic Methods

A series of confirmatory factor analyses (CFA) were conducted using the statistical program Mplus [version 7.31] (Muthén \& Muthén, 2012) for each grade and subject assessment. Mplus is commonly used for collecting validity evidence on the internal structure of assessments. Weighted least squares means and variance adjusted (WLSMV) was employed as the estimation method because it is less sensitive to the size of the sample and the model and is also shown to perform well with categorical variables (Muthén, du Toit, \& Spisic, 1997).

As previously stated, the method of reporting scores used for the State of Florida implies separate factors for each reporting category, connected by a single underlying factor. This model is subsequently referred to as the implied model. In factor analytic terms, this suggests that test items load onto separate first-order factors, with the first-order factors connected to a single underlying second-order factor. The use of the CFA in this section establishes some validity evidence for the degree to which the implied model is reasonable.

A chi-square difference test is often applied to assess model fit by comparing two nested models (i.e., the null model and the alternative model). However, it is sensitive to sample size, almost always rejecting the null hypothesis when the sample size is large. Therefore, instead of conducting a chi-square difference test, other goodness-of-fit indices were used to evaluate the implied model for FSA.

If the internal structure of the test was strictly unidimensional, then the overall person ability measure, theta $(\theta)$, would be the single common factor, and the correlation matrix among test items would suggest no discernable pattern among factors. As such, there would be no empirical or logical basis to report scores for the separate performance categories. In factor analytic terms, a test structure that is strictly unidimensional implies a single-order factor model, in which all test items load onto a single underlying factor. The development below expands the first-order model to a generalized second-order parameterization to show the relationship between the models.

The factor analysis models are based on the matrix $\boldsymbol{S}$ of tetrachoric and polychoric sample correlations among the item scores (Olsson, 1979), and the matrix $\boldsymbol{W}$ of asymptotic covariances among these sample correlations (Jöreskog, 1994) is employed as a weight matrix in a weighted least squares estimation approach (Browne, 1984; Muthén, 1984) to minimize the fit function:

$$
F_{W L S}=\operatorname{vech}(\boldsymbol{S}-\widehat{\boldsymbol{\Sigma}})^{\prime} \boldsymbol{W}^{-\mathbf{1}} \operatorname{vech}(\boldsymbol{S}-\widehat{\boldsymbol{\Sigma}})
$$

In the equation above, $\widehat{\boldsymbol{\Sigma}}$ is the implied correlation matrix, given the estimated factor model, and the function vech vectorizes a symmetric matrix. That is, vech stacks each column of the matrix to form a vector. Note that the WLSMV approach (Muthén, du Toit, \& Spisic, 1997) employs a
weight matrix of asymptotic variances (i.e., the diagonal of the weight matrix) instead of the full asymptotic covariances.

We posit a first-order factor analysis where all test items load onto a single common factor, as the base model. The first-order model can be mathematically represented as:

$$
\widehat{\boldsymbol{\Sigma}}=\boldsymbol{\Lambda} \boldsymbol{\Phi} \boldsymbol{\Lambda}^{\prime}+\boldsymbol{\Theta}
$$

where $\boldsymbol{\Lambda}$ is the matrix of item factor loadings (with $\boldsymbol{\Lambda}^{\prime}$ representing its transpose), and $\boldsymbol{\Theta}$ is the uniqueness, or measurement error. The matrix $\boldsymbol{\Phi}$ is the correlation among the separate factors. For the base model, items are thought only to load onto a single underlying factor. Hence $\boldsymbol{\Lambda}$ is a $p \times 1$ vector, where $p$ is the number of test items and $\boldsymbol{\Phi}$ is a scalar equal to 1 . Therefore, it is possible to drop the matrix $\boldsymbol{\Phi}$ from the general notation. However, this notation is retained to more easily facilitate comparisons to the implied model, such that it can subsequently be viewed as a special case of the second-order factor analysis.

For the implied model, we posit a second-order factor analysis in which test items are coerced to load onto the reporting categories they are designed to target, and all reporting categories share a common underlying factor. The second-order factor analysis can be mathematically represented as:

$$
\widehat{\mathbf{\Sigma}}=\boldsymbol{\Lambda}\left(\boldsymbol{\Gamma} \boldsymbol{\Gamma} \Gamma^{\prime}+\boldsymbol{\Psi}\right) \boldsymbol{\Lambda}^{\prime}+\boldsymbol{\Theta}
$$

where $\hat{\mathbf{\Sigma}}$ is the implied correlation matrix among test items, $\boldsymbol{\Lambda}$ is the $p x k$ matrix of first-order factor loadings relating item scores to first-order factors $\boldsymbol{\Gamma}$ is the $k x l$ matrix of second-order factor loadings relating the first-order factors to the second-order factor with $k$ denoting the number of factors, $\boldsymbol{\Phi}$ is the correlation matrix of the second-order factors, and $\boldsymbol{\Psi}$ is the matrix of first-order factor residuals. All other notation is the same as the first-order model. Note that the second-order model expands the first-order model such that $\boldsymbol{\Phi} \rightarrow \boldsymbol{\Gamma} \boldsymbol{\Phi} \boldsymbol{\Gamma}^{\prime}+\boldsymbol{\Psi}$. As such, the first-order model is said to be nested within the second-order model.

There is a separate factor for each of 4-5 reporting categories for ELA, 3-5 categories for Mathematics, and 3 categories for EOC (see Table 29, Table 30, and Table 31 for reporting category information). Therefore, the number of rows in $\boldsymbol{\Gamma}(k)$ differs between subjects, but the general structure of the factor analysis is consistent across ELA and Mathematics.

The second-order factor model can also be represented graphically, and a sample of the generalized approaches is provided on the following page. The general structure of the second-order factor analysis for ELA is illustrated in Figure 6. This figure is generally representative of the factor analyses performed for all grades and subjects, with the understanding that the number of items within each reporting category could vary across the grades.

The purpose of conducting confirmatory factor analysis for the FSA was to provide evidence that each individual assessment in the FSA implied a second-order factor model: a single underlying second-order factor with the first-order factors defining each of the reporting categories.

Figure 6: Second-Order Factor Model (ELA)

## Generalized Second Order Factor Structure



### 5.2.2 Results

Several goodness-of-fit statistics from each of the analyses are presented in the tables below. Table 46 presents the summary results obtained from confirmatory factor analysis. Three goodness-offit indices were used to evaluate model fit of the item parameters to the manner in which students actually responded to the items. The root mean square error of approximation (RMSEA) is referred to as a badness-of-fit index so that a value closer to 0 implies better fit and a value of 0 implies best fit. In general, RMSEA below 0.05 is considered as good fit and RMSEA over 0.1 suggests poor fit (Browne \& Cudeck, 1993). The Tucker-Lewis index (TLI) and the comparative fit index
(CFI) are incremental goodness-of-fit indices. These indices compare the implied model to the baseline model where no observed variables are correlated (i.e., there are no factors). Values greater than 0.9 are recognized as acceptable, and values over 0.95 are considered as good fit (Hu \& Bentler, 1999). As Hu and Bentler (1999) suggest, the selected cut-off values of the fit index should not be overgeneralized and should be interpreted with caution.

Based on the fit indices, the model showed good fit across content domains. For all tests, RMSEA was below 0.05 , and CFI and TLI were equal to or greater than 0.95 with the exception of grade 8 Mathematics and Algebra 1, which had a CIF and TLI of 0.94.

Table 46: Goodness-of-Fit Second-Order CFA

| ELA |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Grade | df | RMSEA | CFI | TLI | Convergence |
| 3 | 1172 | 0.02 | 0.98 | 0.98 | Yes |
| 4 | 1320 | 0.02 | 0.98 | 0.98 | Yes |
| 5* | 1321 | 0.02 | 0.99 | 0.99 | Yes |
| 6 | 1425 | 0.02 | 0.99 | 0.99 | Yes |
| 7 | 1425 | 0.02 | 0.99 | 0.99 | Yes |
| 8 | 1425 | 0.02 | 0.99 | 0.99 | Yes |
| 9* | 1536 | 0.02 | 0.99 | 0.98 | Yes |
| 10* | 1535 | 0.02 | 0.99 | 0.99 | Yes |
| Mathematics |  |  |  |  |  |
| Grade | df | RMSEA | CFI | TLI | Convergence |
| 3* | 1375 | 0.03 | 0.98 | 0.98 | Yes |
| 4 | 1373 | 0.02 | 0.98 | 0.98 | Yes |
| 5 | 1374 | 0.03 | 0.98 | 0.98 | Yes |
| 6 | 1479 | 0.03 | 0.97 | 0.97 | Yes |
| 7 | 1479 | 0.03 | 0.97 | 0.97 | Yes |
| 8 | 1480 | 0.03 | 0.94 | 0.94 | Yes |
| EOC |  |  |  |  |  |
| Subject/Form | df | RMSEA | CFI | TLI | Convergence |
| Alg 1/Core 12* | 1594 | 0.03 | 0.94 | 0.94 | Yes |
| Alg 1/Core 13* | 1594 | 0.03 | 0.95 | 0.94 | Yes |
| Alg 1/Core 14* | 1593 | 0.03 | 0.97 | 0.97 | Yes |
| Alg 1/Core 15 | 1592 | 0.03 | 0.97 | 0.97 | Yes |
| Geo/Core 9 | 1592 | 0.03 | 0.98 | 0.97 | Yes |
| Geo/Core 10 | 1592 | 0.02 | 0.98 | 0.98 | Yes |
| Geo/Core 11 | 1592 | 0.03 | 0.98 | 0.98 | Yes |

*For these tests, the second-order model was run by constraining the residual variance of a certain factor to zero due to non-significant negative residual variance.

The second-order factor model converged for all tests. However, the residual variance for one factor fell slightly below the boundary of 0 for grades 5, 9, and 10 ELA, grade 3 Mathematics, and Algebra 1 when using the M-Plus software package. This negative residual variance may be related to the computational implementation of the optimization approach in M-Plus, it may be a flag related to model misspecification, or it may be related to other causes (Van Driel, 1978; Chen, Bollen, Paxton, Curran \& Kirby, 2001). The residual variance was constrained to 0 for these tests. This is equivalent to treating the parameter as fixed, which does not necessarily conform to our apriori hypothesis.

As indicated in Section 3.1, FSA items are operationally calibrated by IRTPRO software; however, factor analyses presented here were conducted with Mplus software. There are some noted differences between these software packages in terms of their model parameter estimation algorithms and item-specific measurement models. First, IRTPRO employs full information maximum likelihood and chooses model parameter estimates so that the likelihood of data can be maximized, whereas Mplus uses WLSMV based on limited information maximum likelihood and chooses model parameter estimates so that the likelihood of the observed covariations among items can be maximized. Secondly, IRTPRO allows one to model pseudo-guessing via the 3PL model, whereas Mplus does not include the same flexibility. However, CFA results presented here still indicated good fit indices even though pseudo-guessing was constrained to zero or not taken into account.

In Table 47, Table 48, and Table 49, we provide the estimated correlations between the reporting categories from the second-order factor model for ELA, Mathematics, and EOC, respectively. In all cases, these correlations are very high. However, the results provide empirical evidence that there is some detectable dimensionality among reporting categories.

Table 47: Correlations Among ELA Factors

| Grade | Reporting Category | Cat1 | Cat2 | Cat3 | Cat4 | Cat5 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 3 | Key Ideas and Details (Cat1) | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 0.99 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 0.97 | 0.98 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 0.85 | 0.86 | 0.85 | 1.00 |  |
| 4 | Key Ideas and Details (Cat1) | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 0.96 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 0.94 | 0.96 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 0.90 | 0.91 | 0.90 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 0.66 | 0.67 | 0.66 | 0.63 | 1.00 |
| 5* | Key Ideas and Details (Cat1) | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 0.98 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 0.99 | 0.99 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 0.90 | 0.90 | 0.91 | 1.00 |  |


|  | Text-Based Writing (Cat5) | 0.69 | 0.70 | 0.70 | 0.63 | 1.00 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 6 | Key Ideas and Details (Cat1) | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 0.99 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 0.98 | 0.98 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 0.92 | 0.92 | 0.92 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 0.68 | 0.68 | 0.68 | 0.64 | 1.00 |
| 7 | Key Ideas and Details (Cat1) | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 0.98 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 0.99 | 0.98 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 0.86 | 0.85 | 0.86 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 0.71 | 0.71 | 0.71 | 0.62 | 1.00 |
| 8 | Key Ideas and Details (Cat1) | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 0.99 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 0.96 | 0.97 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 0.93 | 0.94 | 0.91 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 0.75 | 0.75 | 0.73 | 0.71 | 1.00 |
| 9* | Key Ideas and Details (Cat1) | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 1.00 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 0.995 | 0.995 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 0.77 | 0.77 | 0.76 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 0.75 | 0.75 | 0.75 | 0.58 | 1.00 |
| 10* | Key Ideas and Details (Cat1) | 1.00 |  |  |  |  |
|  | Craft and Structure (Cat2) | 0.998 | 1.00 |  |  |  |
|  | Integration of Knowledge and Ideas (Cat3) | 0.97 | 0.97 | 1.00 |  |  |
|  | Language and Editing Task (Cat4) | 0.90 | 0.89 | 0.87 | 1.00 |  |
|  | Text-Based Writing (Cat5) | 0.79 | 0.79 | 0.76 | 0.71 | 1.00 |

*For these tests, the second-order model was run by constraining the residual variance of a certain factor to zero due to non-significant negative residual variance.

Table 48: Correlations Among Mathematics Factors

| Grade | Reporting Category | Cat1 | Cat2 | Cat3 | Cat4 | Cat5 |
| :--- | :--- | :--- | :--- | :--- | :--- | :---: |
| $3^{*}$ | Operations, Algebraic Thinking, <br> and Numbers in Base Ten (Cat1) | 1.00 |  |  |  |  |


|  | Numbers and Operations Fractions (Cat2) | 0.90 | 1.00 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Measurement, Data, and Geometry (Cat3) | 0.97 | 0.93 | 1.00 |  |  |
| 4 | Operations and Algebraic Thinking (Cat1) | 1.00 |  |  |  |  |
|  | Numbers and Operations in Base Ten (Cat2) | 0.96 | 1.00 |  |  |  |
|  | Numbers and Operations Fractions (Cat3) | 0.99 | 0.96 | 1.00 |  |  |
|  | Measurement, Data, and Geometry (Cat4) | 0.97 | 0.94 | 0.97 | 1.00 |  |
| 5 | Operations, Algebraic Thinking, and Fractions (Cat1) | 1.00 |  |  |  |  |
|  | Numbers and Operations in Base Ten (Cat2) | 0.96 | 1.00 |  |  |  |
|  | Measurement, Data, and Geometry (Cat3) | 0.96 | 0.95 | 1.00 |  |  |
| 6 | Ratio and Proportional Relationships (Cat1) | 1.00 |  |  |  |  |
|  | Expressions and Equations (Cat2) | 0.98 | 1.00 |  |  |  |
|  | Geometry (Cat3) | 0.96 | 0.96 | 1.00 |  |  |
|  | Statistics and Probability (Cat4) | 0.95 | 0.95 | 0.93 | 1.00 |  |
|  | The Number System (Cat5) | 0.99 | 0.99 | 0.97 | 0.96 | 1.00 |
| 7 | Ratio and Proportional Relationships (Cat1) | 1.00 |  |  |  |  |
|  | Expressions and Equations (Cat2) | 0.96 | 1.00 |  |  |  |
|  | Geometry (Cat3) | 0.91 | 0.92 | 1.00 |  |  |
|  | Statistics and Probability (Cat4) | 0.95 | 0.97 | 0.91 | 1.00 |  |
|  | The Number System (Cat5) | 0.97 | 0.98 | 0.93 | 0.97 | 1.00 |
| 8 | Expressions and Equations (Cat1) | 1.00 |  |  |  |  |
|  | Functions (Cat2) | 0.93 | 1.00 |  |  |  |
|  | Geometry (Cat3) | 0.91 | 0.88 | 1.00 |  |  |
|  | Statistics and Probability and the Number System (Cat4) | 0.95 | 0.92 | 0.90 | 1.00 |  |

*For these tests, the second-order model was run by constraining the residual variance of a certain factor to zero due to non-significant negative residual variance.

Table 49: Correlations Among EOC Factors

| Course/Form | Reporting Category | Cat1 | Cat2 | Cat3 |
| :--- | :--- | :--- | :---: | :---: |
| Algebra 1/Core 12* | Algebra and Modeling (Cat1) | 1.00 |  |  |


|  | Functions and Modeling (Cat2) | 0.97 | 1.00 |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Statistics and the Number System (Cat3) | 0.99 | 0.95 | 1.00 |
| Algebra 1/Core 13* | Algebra and Modeling (Cat1) | 1.00 |  |  |
|  | Functions and Modeling (Cat2) | 0.97 | 1.00 |  |
|  | Statistics and the Number System (Cat3) | 0.98 | 0.95 | 1.00 |
| Algebra 1/Core 14* | Algebra and Modeling (Cat1) | 1.00 |  |  |
|  | Functions and Modeling (Cat2) | 0.98 | 1.00 |  |
|  | Statistics and the Number System (Cat3) | 0.96 | 0.98 | 1.00 |
| Algebra 1/Core 15 | Algebra and Modeling (Cat1) | 1.00 |  |  |
|  | Functions and Modeling (Cat2) | 0.97 | 1.00 |  |
|  | Statistics and the Number System (Cat3) | 0.95 | 0.97 | 1.00 |
| Geometry/Core 9 | Congruence, Similarity, Right Triangles, and Trigonometry (Cat1) | 1.00 |  |  |
|  | Circles, Geometric Measurement, and Geometric Properties with Equations (Cat2) | 0.98 | 1.00 |  |
|  | Modeling with Geometry (Cat3) | 0.97 | 0.96 | 1.00 |
| Geometry/Core 10 | Congruence, Similarity, Right Triangles, and Trigonometry (Cat1) | 1.00 |  |  |
|  | Circles, Geometric Measurement, and Geometric Properties with Equations (Cat2) | 0.98 | 1.00 |  |
|  | Modeling with Geometry (Cat3) | 0.97 | 0.97 | 1.00 |
| Geometry/Core 11 | Congruence, Similarity, Right Triangles, and Trigonometry (Cat1) | 1.00 |  |  |
|  | Circles, Geometric Measurement, and Geometric Properties with Equations (Cat2) | 0.98 | 1.00 |  |
|  | Modeling with Geometry (Cat3) | 0.97 | 0.96 | 1.00 |

*For these tests, the second-order model was run by constraining the residual variance of a certain factor to zero due to non-significant negative residual variance.

### 5.2.3 Discussion

In all scenarios, the empirical results suggest the implied model fits the data well. That is, these results indicate that reporting an overall score in addition to separate scores for the individual reporting categories is reasonable, as the intercorrelations among items suggest that there are detectable distinctions among reporting categories.

Clearly, the correlations among the separate factors are high, which is reasonable. This again provides support for the measurement model, given that the calibration of all items is performed concurrently. If the correlations among factors were very low, this could possibly suggest that a different IRT model would be needed (e.g., multidimensional IRT) or that the IRT calibration should be performed separately for items measuring different factors. The high correlations among
the factors suggest these alternative methods are unnecessary and that our current approach is in fact preferable.

Overall, these results provide empirical evidence and justification for the use of our scoring and reporting methods. Additionally, the results provide justification for the current IRT model employed.

### 5.3 Local Independence

The validity of the application of IRT depends greatly on meeting the underlying assumptions of the models. One such assumption is local independence, which means that for a given proficiency estimate, the (marginal) likelihood is maximized, assuming the probability of correct responses is the product of independent probabilities over all items (Chen \& Thissen, 1997):

$$
\mathrm{L}(\theta)=\int \prod_{j=1}^{K} \operatorname{Pr}\left(x_{j} \mid \theta\right) f(\theta) \mathrm{d} \theta
$$

When local independence is not met, there are issues of multidimensionality that are unaccounted for in the modeling of the data (Bejar, 1980). In fact, Lord (1980) noted that "local independence follows automatically from unidimensionality" (as cited in Bejar, 1980, p. 5). From a dimensionality perspective, there may be nuisance factors that are influencing relationships among certain items, after accounting for the intended construct of interest. These nuisance factors can be influenced by a number of testing features, such as speededness, fatigue, item chaining, and item or response formats (Yen, 1993).

Yen's $Q_{3}$ statistic (Yen, 1984) was used to measure local independence, which was derived from the correlation between the performances of two items. Simply, the $\mathrm{Q}_{3}$ statistic is the correlation among IRT residuals and is computed using the following equations:

$$
d_{i j}=u_{i j}-T_{j}\left(\widehat{\theta}_{i}\right)
$$

where $u_{i j}$ is the item score of the $i$ th examinee for item $j, T_{j}\left(\hat{\theta}_{i}\right)$ is the estimated true score for item $j$ of examinee $i$, which is defined as

$$
T_{j}\left(\hat{\theta}_{i}\right)=\sum_{k=1}^{m} y_{j k} P_{j k}\left(\hat{\theta}_{i}\right)
$$

where $y_{j k}$ is the weight for response category $k, m$ is the number of response categories, and $P_{j k}\left(\hat{\theta}_{i}\right)$ is the probability of response category $k$ to item $j$ by examinee $i$ with the ability estimate $\hat{\theta}_{i}$.

The pairwise index of local dependence $\mathrm{Q}_{3}$ between item $j$ and item $j$ ' is

$$
Q_{3 j j^{\prime}}=r\left(d_{j}, d_{j^{\prime}}\right)
$$

where $r$ refers to the Pearson product-moment correlation.

When there are $n$ items, $n(n-1) / 2, \mathrm{Q}_{3}$ statistics will be produced. The $\mathrm{Q}_{3}$ values are expected to be small. Table 50, Table 51, and Table 52 present summaries of the distributions of $Q_{3}$ statisticsminimum, 5th percentile, median, 95 th percentile, and maximum values from each grade and subject. The results show that at least $90 \%$ of the items, between the 5 th and 95 th percentiles, for all grades and subjects were smaller than a critical value of 0.2 for $\left|Q_{3}\right|$ (Chen \& Thissen, 1997).

Table 50: ELA Q 3 Statistic

| Grade | Unconditional <br> Observed <br> Correlation | Q3 Distribution |  |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  |  | 5th <br> Percentile | Median | 95th <br> Percentile | Maximum | Minimum | Maximum |  |
| 3 |  | -0.077 | -0.046 | -0.020 | 0.015 | 0.241 | -0.057 | 0.149 |
| 4 | 0.229 | -0.114 | -0.057 | -0.014 | 0.012 | 0.697 | -0.064 | 0.066 |
| 5 | 0.293 | -0.142 | -0.061 | -0.015 | 0.013 | 0.686 | -0.076 | 0.090 |
| 6 | 0.276 | -0.106 | -0.060 | -0.015 | 0.010 | 0.738 | -0.070 | 0.112 |
| 7 | 0.233 | -0.132 | -0.061 | -0.012 | 0.010 | 0.802 | -0.051 | 0.115 |
| 8 | 0.291 | -0.141 | -0.065 | -0.012 | 0.011 | 0.818 | -0.037 | 0.097 |
| 9 | 0.270 | -0.127 | -0.064 | -0.011 | 0.012 | 0.772 | -0.045 | 0.201 |
| 10 | 0.268 | -0.160 | -0.071 | -0.010 | 0.016 | 0.884 | -0.028 | 0.160 |

* Maximum $\mathrm{Q}_{3}$ values of grades 4 through 10 are from elaboration and organization dimensions of the Writing prompt.
** Within Passage $\mathrm{Q}_{3}$ values are computed for each item pair within a passage.
Table 51: Mathematics Q3 Statistic

| Grade | Unconditional <br> Observed <br> Correlation | Minimum |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  |  | 5th <br> Percentile | Median | 95th <br> Percentile | Maximum |  |
| 3 |  | -0.115 | -0.055 | -0.017 | 0.027 | 0.277 |
| 4 | 0.416 | -0.115 | -0.047 | -0.018 | 0.021 | 0.199 |
| 5 | 0.432 | -0.125 | -0.054 | -0.019 | 0.018 | 0.255 |
| 6 | 0.409 | -0.084 | -0.047 | -0.017 | 0.018 | 0.222 |
| 7 | 0.385 | -0.117 | -0.056 | -0.015 | 0.020 | 0.399 |
| 8 | 0.291 | -0.111 | -0.052 | -0.016 | 0.021 | 0.404 |

Table 52: EOC Q3 Statistic

| Course | Unconditional <br> Observed <br> Correlation* | Minimum |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  |  | 5th <br> Percentile | Median | 95th <br> Percentile | Maximum |  |
| Algebra 1 |  | -0.119 | -0.049 | -0.013 | 0.013 | 0.271 |
| Geometry |  | -0.388 | -0.151 | -0.014 | 0.120 | 0.401 |

* Unconditional observed correlations were computed based on each core form and averaged over those core forms.


### 5.4 Convergent and Discriminant Validity

According to Standard 1.14 of the Standards for Educational and Psychological Testing (AERA, APA, and NCME, 1999), it is necessary to provide evidence of convergent and discriminant validity evidence. It is a part of validity evidence demonstrating that assessment scores are related as expected with criterion and other variables for all student groups. However, a second, independent test measuring the same constructs as Mathematics and ELA in the State of Florida, which could easily permit for a cross test set of correlations, was not available. Therefore, the correlations between subscores within and across Mathematics and ELA were examined alternatively. The a-priori expectation is that subscores within the same subject (e.g., Mathematics) will correlate more positively than subscore correlations across subjects (e.g., Mathematics and ELA). These correlations are based on a small number of items (e.g., typically around 8 to 12); as a consequence, the observed score correlations will be smaller in magnitude as a result of the very large measurement error at the subscore level. For this reason, both the observed score and the disattenuated correlations are provided.

Observed and disattenuated subscore correlations were calculated both within subject and across subjects for grades 3-8 Mathematics and ELA. In general, the pattern is consistent with the a-priori expectation that subscores within a test correlate more highly than correlations between tests measuring a different construct with a few small notes on the writing dimensions. The writing dimensions are based on a single essay, which is scored as three test items. Hence, the correlations between writing and other dimensions, both in the observed score and even disattenuated scores, are somewhat unstable given the large measurement error. Tables 53-64 show the observed and disattenuated score correlations between Mathematics and ELA subscores for grades 3-8, where students took both subjects.

Table 53: Grade 3 Observed Score Correlations

| Subject | Reporting Category | Mathematics |  |  | ELA |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Rep1 | Rep2 | Rep3 | Rep1 | Rep2 | Rep3 | Rep4 |
| Mathematics | Operations, Algebraic Thinking, and Numbers in Base Ten (Cat1) | 1.00 | 0.72 | 0.83 | 0.63 | 0.64 | 0.54 | 0.57 |
|  | Numbers and Operations - Fractions (Cat2) |  | 1.00 | 0.74 | 0.57 | 0.57 | 0.51 | 0.50 |
|  | Measurement, Data, and Geometry (Cat3) |  |  | 1.00 | 0.67 | 0.68 | 0.60 | 0.58 |
| ELA | Key Ideas and Details (Cat1) |  |  |  | 1.00 | 0.77 | 0.65 | 0.61 |
|  | Craft and Structure (Cat2) |  |  |  |  | 1.00 | 0.66 | 0.61 |
|  | Integration of Knowledge and Ideas (Cat3) |  |  |  |  |  | 1.00 | 0.51 |
|  | Language and Editing Task (Cat4) |  |  |  |  |  |  | 1.00 |

Table 54: Grade 3 Disattenuated Score Correlations

|  | Reporting Category | Mathematics |  |  | ELA |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Rep1 | Rep2 | Rep3 | Rep1 | Rep2 | Rep3 | Rep4 |
| Mathematics | Operations, Algebraic Thinking, and Numbers in Base Ten (Cat1) | 1.00 | 0.89 | 0.95 | 0.75 | 0.76 | 0.74 | 0.74 |
|  | Numbers and Operations - Fractions (Cat2) |  | 1.00 | 0.96 | 0.76 | 0.77 | 0.80 | 0.73 |
|  | Measurement, Data, and Geometry (Cat3) |  |  | 1.00 | 0.83 | 0.84 | 0.86 | 0.77 |
| ELA | Key Ideas and Details (Cat1) |  |  |  | 1.00 | 0.99 | 0.97 | 0.85 |
|  | Craft and Structure (Cat2) |  |  |  |  | 1.00 | 0.98 | 0.86 |
|  | Integration of Knowledge and Ideas (Cat3) |  |  |  |  |  | 1.00 | 0.83 |
|  | Language and Editing Task (Cat4) |  |  |  |  |  |  | 1.00 |

Table 55: Grade 4 Observed Score Correlations

| Subject | Reporting Category | Mathematics |  |  |  | ELA |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Rep1 | Rep2 | Rep3 | Rep4 | Rep1 | Rep2 | Rep3 | Rep4 | Rep5 |
| Mathematics | Operations and Algebraic Thinking (Cat1) | 1.00 | 0.74 | 0.77 | 0.76 | 0.56 | 0.56 | 0.50 | 0.55 | 0.51 |
|  | Numbers and Operations in Base Ten (Cat2) |  | 1.00 | 0.79 | 0.78 | 0.54 | 0.55 | 0.49 | 0.54 | 0.51 |
|  | Numbers and Operations - Fractions (Cat3) |  |  | 1.00 | 0.83 | 0.59 | 0.59 | 0.53 | 0.58 | 0.53 |
|  | Measurement, Data, and Geometry (Cat4) |  |  |  | 1.00 | 0.59 | 0.59 | 0.53 | 0.58 | 0.53 |
| ELA | Key Ideas and Details (Cat1) |  |  |  |  | 1.00 | 0.67 | 0.68 | 0.60 | 0.50 |
|  | Craft and Structure (Cat2) |  |  |  |  |  | 1.00 | 0.61 | 0.57 | 0.49 |
|  | Integration of Knowledge and Ideas (Cat3) |  |  |  |  |  |  | 1.00 | 0.55 | 0.45 |
|  | Language and Editing Task (Cat4) |  |  |  |  |  |  |  | 1.00 | 0.51 |
|  | Text-Based Writing (Cat5) |  |  |  |  |  |  |  |  | 1.00 |

Table 56: Grade 4 Disattenuated Score Correlations

| Subject | Reporting Category | Mathematics |  |  |  | ELA |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Rep1 | Rep2 | Rep3 | Rep4 | Rep1 | Rep2 | Rep3 | Rep4 | Rep5 |
| Mathematics | Operations and Algebraic Thinking (Cat1) | 1.00 | 0.98 | 0.99 | 0.97 | 0.76 | 0.82 | 0.75 | 0.81 | 0.66 |
|  | Numbers and Operations in Base Ten (Cat2) |  | 1.00 | 0.97 | 0.94 | 0.71 | 0.76 | 0.69 | 0.77 | 0.63 |
|  | Numbers and Operations - Fractions (Cat3) |  |  | 1.00 | 0.97 | 0.75 | 0.80 | 0.74 | 0.79 | 0.64 |
|  | Measurement, Data, and Geometry (Cat4) |  |  |  | 1.00 | 0.73 | 0.79 | 0.72 | 0.78 | 0.63 |
| ELA | Key Ideas and Details (Cat1) |  |  |  |  | 1.00 | 0.95 | 0.99 | 0.87 | 0.64 |
|  | Craft and Structure (Cat2) |  |  |  |  |  | 1.00 | 0.96 | 0.89 | 0.67 |
|  | Integration of Knowledge and Ideas (Cat3) |  |  |  |  |  |  | 1.00 | 0.87 | 0.63 |
|  | Language and Editing Task (Cat4) |  |  |  |  |  |  |  | 1.00 | 0.71 |
|  | Text-Based Writing (Cat5) |  |  |  |  |  |  |  |  | 1.00 |

Table 57: Grade 5 Observed Score Correlations

| Subject | Reporting Category | Mathematics |  |  | ELA |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Rep1 | Rep2 | Rep3 | Rep1 | Rep2 | Rep3 | Rep4 | Rep5 |
| Mathematics | Operations, Algebraic Thinking, and Fractions (Cat1) | 1.00 | 0.85 | 0.84 | 0.65 | 0.61 | 0.63 | 0.56 | 0.56 |
|  | Numbers and Operations in Base Ten (Cat2) |  | 1.00 | 0.81 | 0.61 | 0.57 | 0.59 | 0.53 | 0.53 |
|  | Measurement, Data, and Geometry (Cat3) |  |  | 1.00 | 0.65 | 0.60 | 0.63 | 0.54 | 0.53 |
| ELA | Key Ideas and Details (Cat1) |  |  |  | 1.00 | 0.73 | 0.78 | 0.59 | 0.55 |
|  | Craft and Structure (Cat2) |  |  |  |  | 1.00 | 0.72 | 0.57 | 0.53 |
|  | Integration of Knowledge and Ideas (Cat3) |  |  |  |  |  | 1.00 | 0.58 | 0.54 |
|  | Language and Editing Task (Cat4) |  |  |  |  |  |  | 1.00 | 0.51 |
|  | Text-Based Writing (Cat5) |  |  |  |  |  |  |  | 1.00 |

Table 58: Grade 5 Disattenuated Score Correlations

| Subject | Reporting Category | Mathematics |  |  | ELA |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Rep1 | Rep2 | Rep3 | Rep1 | Rep2 | Rep3 | Rep4 | Rep5 |
| Mathematics | Operations, Algebraic Thinking, and Fractions (Cat1) | 1.00 | 0.96 | 0.97 | 0.78 | 0.77 | 0.78 | 0.79 | 0.65 |
|  | Numbers and Operations in Base Ten (Cat2) |  | 1.00 | 0.94 | 0.73 | 0.73 | 0.73 | 0.75 | 0.62 |
|  | Measurement, Data, and Geometry (Cat3) |  |  | 1.00 | 0.80 | 0.78 | 0.79 | 0.78 | 0.63 |
| ELA | Key Ideas and Details (Cat1) |  |  |  | 1.00 | 0.98 | 1.02 | 0.88 | 0.68 |
|  | Craft and Structure (Cat2) |  |  |  |  | 1.00 | 1.00 | 0.90 | 0.69 |
|  | Integration of Knowledge and Ideas (Cat3) |  |  |  |  |  | 1.00 | 0.89 | 0.69 |
|  | Language and Editing Task (Cat4) |  |  |  |  |  |  | 1.00 | 0.73 |
|  | Text-Based Writing (Cat5) |  |  |  |  |  |  |  | 1.00 |

Table 59: Grade 6 Observed Score Correlations

| Subject | Reporting Category | Mathematics |  |  |  |  | ELA |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Rep1 | Rep2 | Rep3 | Rep4 | Rep5 | Rep1 | Rep2 | Rep3 | Rep4 | Rep5 |
| Mathematics | Ratio and Proportional Relationships (Cat1) | 1.00 | 0.77 | 0.68 | 0.69 | 0.76 | 0.60 | 0.59 | 0.56 | 0.56 | 0.50 |
|  | Expressions and Equations (Cat2) |  | 1.00 | 0.75 | 0.75 | 0.84 | 0.64 | 0.63 | 0.61 | 0.58 | 0.53 |
|  | Geometry (Cat3) |  |  | 1.00 | 0.68 | 0.74 | 0.57 | 0.55 | 0.55 | 0.49 | 0.45 |
|  | Statistics and Probability (Cat4) |  |  |  | 1.00 | 0.73 | 0.62 | 0.62 | 0.59 | 0.56 | 0.50 |
|  | The Number System (Cat5) |  |  |  |  | 1.00 | 0.62 | 0.61 | 0.59 | 0.56 | 0.52 |
| ELA | Key Ideas and Details (Cat1) |  |  |  |  |  | 1.00 | 0.73 | 0.70 | 0.60 | 0.51 |
|  | Craft and Structure (Cat2) |  |  |  |  |  |  | 1.00 | 0.66 | 0.60 | 0.50 |
|  | Integration of Knowledge and Ideas (Cat3) |  |  |  |  |  |  |  | 1.00 | 0.54 | 0.48 |
|  | Language and Editing Task (Cat4) |  |  |  |  |  |  |  |  | 1.00 | 0.53 |
|  | Text-Based Writing (Cat5) |  |  |  |  |  |  |  |  |  | 1.00 |

Table 60: Grade 6 Disattenuated Score Correlations

| Subject | Reporting Category | Mathematics |  |  |  |  | ELA |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Rep1 | Rep2 | Rep3 | Rep4 | Rep5 | Rep1 | Rep2 | Rep3 | Rep4 | Rep5 |
| Mathematics | Ratio and Proportional Relationships (Cat1) | 1.00 | 0.99 | 0.94 | 0.96 | 1.00 | 0.81 | 0.83 | 0.82 | 0.82 | 0.65 |
|  | Expressions and Equations (Cat2) |  | 1.00 | 0.94 | 0.94 | 1.01 | 0.80 | 0.81 | 0.81 | 0.77 | 0.63 |
|  | Geometry (Cat3) |  |  | 1.00 | 0.92 | 0.95 | 0.75 | 0.76 | 0.78 | 0.70 | 0.58 |
|  | Statistics and Probability (Cat4) |  |  |  | 1.00 | 0.94 | 0.83 | 0.85 | 0.83 | 0.80 | 0.64 |
|  | The Number System (Cat5) |  |  |  |  | 1.00 | 0.79 | 0.80 | 0.80 | 0.77 | 0.63 |
| ELA | Key Ideas and Details (Cat1) |  |  |  |  |  | 1.00 | 0.99 | 0.98 | 0.84 | 0.64 |
|  | Craft and Structure (Cat2) |  |  |  |  |  |  | 1.00 | 0.96 | 0.87 | 0.65 |
|  | Integration of Knowledge and Ideas (Cat3) |  |  |  |  |  |  |  | 1.00 | 0.81 | 0.64 |
|  | Language and Editing Task (Cat4) |  |  |  |  |  |  |  |  | 1.00 | 0.72 |
|  | Text-Based Writing (Cat5) |  |  |  |  |  |  |  |  |  | 1.00 |

Table 61: Grade 7 Observed Score Correlations

| Subject | Reporting Category | Mathematics |  |  |  |  | ELA |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Rep1 | Rep2 | Rep3 | Rep4 | Rep5 | Rep1 | Rep2 | Rep3 | Rep4 | Rep5 |
| Mathematics | Ratio and Proportional Relationships (Cat1) | 1.00 | 0.76 | 0.76 | 0.75 | 0.77 | 0.57 | 0.61 | 0.50 | 0.43 | 0.50 |
|  | Expressions and Equations (Cat2) |  | 1.00 | 0.73 | 0.67 | 0.74 | 0.56 | 0.58 | 0.49 | 0.40 | 0.48 |
|  | Geometry (Cat3) |  |  | 1.00 | 0.68 | 0.73 | 0.55 | 0.58 | 0.48 | 0.41 | 0.48 |
|  | Statistics and Probability (Cat4) |  |  |  | 1.00 | 0.69 | 0.56 | 0.59 | 0.49 | 0.41 | 0.48 |
|  | The Number System (Cat5) |  |  |  |  | 1.00 | 0.56 | 0.58 | 0.49 | 0.41 | 0.48 |
| ELA | Key Ideas and Details (Cat1) |  |  |  |  |  | 1.00 | 0.72 | 0.61 | 0.46 | 0.54 |
|  | Craft and Structure (Cat2) |  |  |  |  |  |  | 1.00 | 0.61 | 0.46 | 0.53 |
|  | Integration of Knowledge and Ideas (Cat3) |  |  |  |  |  |  |  | 1.00 | 0.38 | 0.45 |
|  | Language and Editing Task (Cat4) |  |  |  |  |  |  |  |  | 1.00 | 0.41 |
|  | Text-Based Writing (Cat5) |  |  |  |  |  |  |  |  |  | 1.00 |

Table 62: Grade 7 Disattenuated Score Correlations

| Subject | Reporting Category | Mathematics |  |  |  |  | ELA |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Rep1 | Rep2 | Rep3 | Rep4 | Rep5 | Rep1 | Rep2 | Rep3 | Rep4 | Rep5 |
| Mathematics | Ratio and Proportional Relationships (Cat1) | 1.00 | 0.96 | 0.92 | 0.99 | 0.98 | 0.72 | 0.75 | 0.72 | 0.67 | 0.59 |
|  | Expressions and Equations (Cat2) |  | 1.00 | 0.95 | 0.96 | 1.01 | 0.75 | 0.77 | 0.75 | 0.68 | 0.61 |
|  | Geometry (Cat3) |  |  | 1.00 | 0.93 | 0.96 | 0.71 | 0.73 | 0.71 | 0.66 | 0.58 |
|  | Statistics and Probability (Cat4) |  |  |  | 1.00 | 0.98 | 0.80 | 0.82 | 0.78 | 0.73 | 0.64 |
|  | The Number System (Cat5) |  |  |  |  | 1.00 | 0.76 | 0.78 | 0.76 | 0.70 | 0.61 |
| ELA | Key Ideas and Details (Cat1) |  |  |  |  |  | 1.00 | 0.95 | 0.93 | 0.78 | 0.68 |
|  | Craft and Structure (Cat2) |  |  |  |  |  |  | 1.00 | 0.92 | 0.76 | 0.66 |
|  | Integration of Knowledge and Ideas (Cat3) |  |  |  |  |  |  |  | 1.00 | 0.73 | 0.65 |
|  | Language and Editing Task (Cat4) |  |  |  |  |  |  |  |  | 1.00 | 0.65 |
|  | Text-Based Writing (Cat5) |  |  |  |  |  |  |  |  |  | 1.00 |

Table 63: Grade 8 Observed Score Correlations

| Subject | Reporting Category | Mathematics |  |  |  | ELA |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Rep1 | Rep2 | Rep3 | Rep4 | Rep1 | Rep2 | Rep3 | Rep4 | Rep5 |
| Mathematics | Expressions and Equations (Cat1) | 1.00 | 0.70 | 0.70 | 0.71 | 0.45 | 0.51 | 0.44 | 0.39 | 0.44 |
|  | Functions (Cat2) |  | 1.00 | 0.66 | 0.65 | 0.46 | 0.51 | 0.46 | 0.38 | 0.43 |
|  | Geometry (Cat3) |  |  | 1.00 | 0.65 | 0.43 | 0.48 | 0.42 | 0.36 | 0.40 |
|  | Statistics and Probability and the Number System (Cat4) |  |  |  | 1.00 | 0.44 | 0.50 | 0.44 | 0.40 | 0.45 |
| ELA | Key Ideas and Details (Cat1) |  |  |  |  | 1.00 | 0.69 | 0.61 | 0.46 | 0.47 |
|  | Craft and Structure (Cat2) |  |  |  |  |  | 1.00 | 0.65 | 0.54 | 0.55 |
|  | Integration of Knowledge and Ideas (Cat3) |  |  |  |  |  |  | 1.00 | 0.45 | 0.46 |
|  | Language and Editing Task (Cat4) |  |  |  |  |  |  |  | 1.00 | 0.50 |
|  | Text-Based Writing (Cat5) |  |  |  |  |  |  |  |  | 1.00 |

Table 64: Grade 8 Disattenuated Score Correlations

| Subject | Reporting Category | Mathematics |  |  |  | ELA |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Rep1 | Rep2 | Rep3 | Rep4 | Rep1 | Rep2 | Rep3 | Rep4 | Rep5 |
| Mathematics | Expressions and Equations (Cat1) | 1.00 | 0.94 | 0.92 | 0.96 | 0.60 | 0.63 | 0.61 | 0.58 | 0.53 |
|  | Functions (Cat2) |  | 1.00 | 0.92 | 0.93 | 0.66 | 0.68 | 0.67 | 0.61 | 0.56 |
|  | Geometry (Cat3) |  |  | 1.00 | 0.91 | 0.60 | 0.62 | 0.60 | 0.57 | 0.50 |
|  | Statistics and Probability and the Number System (Cat4) |  |  |  | 1.00 | 0.63 | 0.67 | 0.64 | 0.64 | 0.57 |
| ELA | Key Ideas and Details (Cat1) |  |  |  |  | 1.00 | 0.91 | 0.88 | 0.74 | 0.61 |
|  | Craft and Structure (Cat2) |  |  |  |  |  | 1.00 | 0.89 | 0.80 | 0.66 |
|  | Integration of Knowledge and Ideas (Cat3) |  |  |  |  |  |  | 1.00 | 0.73 | 0.61 |
|  | Language and Editing Task (Cat4) |  |  |  |  |  |  |  | 1.00 | 0.73 |
|  | Text-Based Writing (Cat5) |  |  |  |  |  |  |  |  | 1.00 |

## 6. Evidence of Comparability

As the FSA was administered in multiple modes (both online and paper-pencil), it is important to provide evidence of comparability between the versions. If the content between forms varies, then one cannot justify score comparability.

Student scores should not depend on the mode of administration or the type of test form. FSA had online assessments for grades 4 through 10 ELA, grades 3 through 8 Mathematics, and EOC. To improve the accessibility of the statewide assessment, alternate assessments were provided to students whose Individual Educational Plans (IEP) or Section 504 Plans indicated such a need. Thus, the comparability of scores obtained via alternate means of administration must be established and evaluated. For grade 3 Reading, there was no accommodated form, as these tests were universally administered on paper. For other grades, the number of items replaced between the online and paper accommodated forms is provided in Table 65. In EOC, the first core form (Core 12 for Algebra 1 and Core 9 Geometry) was administered as the accommodated version.

### 6.1 Match-with-Test Blueprints for Both Paper-Pencil and Online Tests

For the 2017-2018 FSA, the paper-pencil versions of the tests were developed according to the same test specifications used for the online tests. These paper tests matched the same blueprints designed for the online tests. In this section, evidence of matching blueprints for both online and paper tests is provided. The procedures used to establish comparable forms are provided in Volume 2, Test Development, of the 2018 FSA Technical Report.

### 6.2 Comparability of FSA Test Scores Over Time

The comparability of FSA scores over time was ensured via two methods. First, during test construction, both a content and statistical perspective were implemented. The FSA test items placed onto forms in spring 2015, spring 2016, spring 2017, and spring 2018 were aligned to the same standards and test blueprint specifications. In addition, spring 2017 form statistics were used as targets for both numerical and graphical summaries for the spring 2018 forms. See Section 4 of Volume 2 for details about both the content and statistical methods. Second, during spring 2018 calibrations, equating was performed in order to place item parameters estimates from spring 2018 onto the 2015 baseline scale. The equating procedure and results are presented in Volume 1 Section 6.2.

### 6.3 COMPARABILITY OF ONLINE AND PAPER-PENCIL TEST SCORES

Test forms for paper-pencil administration were offered as a special accommodation for students who qualified, according to their Individual Educational Plans (IEP) or Section 504 Plans. These forms aligned to the same test specifications as the online forms and used the same item parameters for scoring for items that were common in both forms. However, without an online system, technology-enhanced items could not be administered with paper-pencil testing. Thus, some items were replaced with comparable items formatted for paper. This was the only difference between the two versions.

After replacing technology-enhanced items with multiple-choice items, accommodated forms were somewhat different from online forms. This pattern can be easily found in test characteristic curves (TCCs) for Mathematics, in which a relatively large number of items were replaced in the accommodated forms. However, this is not concerning since all of the items are on the same IRT scale. In EOC, TCCs for the accommodated forms are above those of the online forms, slightly shifted upward compared to the TCCs for the online forms. Conversely, there is an overlap of TCCs for the online and accommodated forms in ELA. As seen in Table 65, all ELA accommodated form items were the same as online form items. TCCs for Mathematics are presented in Appendix E.

Table 65: Number of Item Replacements for the Accommodated Forms

| Mathematics | ELA | Number of Items Replaced | Items Replaced |
| :--- | :--- | :--- | :--- |
| Grade 3 | 5 | Grade 4 | 0 |
| Grade 4 | 8 | Grade 5 | 0 |
| Grade 5 | 7 | Grade 6 | 0 |
| Grade 6 | 19 | Grade 7 | 0 |
| Grade 7 | 18 | Grade 8 | 0 |
| Grade 8 | 13 | Grade 9 | 0 |
| Algebra 1 | 8 | Grade 10 | 0 |
| Geometry | 10 |  |  |

A device comparability study was conducted to provide evidence of the comparability of the FSA tests across the most frequently used platforms. Score comparability across different devices can be examined to assess whether student performance on the FSA differs between students conditional on the device. The device effects were examined via regression and a likelihood ratio test to compare the regression models. The study showed that there are no systematic differences in the scores for students when administered the FSA on different devices. The details of the study can be found in Appendix F.

## 7. Fairness and Accessibility

### 7.1 Fairness in Content

The principles of universal design of assessments provide guidelines for test design to minimize the impact of construct-irrelevant factors in assessing student achievement. Universal design removes barriers to provide access for the widest range of students possible. Seven principles of universal design are applied in the process of test development (Thompson, Johnstone, \& Thurlow, 2002):

1. Inclusive assessment population
2. Precisely defined constructs
3. Accessible, non-biased items
4. Amenable to accommodations
5. Simple, clear, and intuitive instructions and procedures
6. Maximum readability and comprehensibility
7. Maximum legibility

Test development specialists have received extensive training on the principles of universal design and apply these principles in the development of all test materials. In the review process, adherence to the principles of universal design is verified by Florida educators and stakeholders.

### 7.2 Statistical Fairness in Item Statistics

Analysis of the content alone is not sufficient to determine the fairness of a test. Rather, it must be accompanied by statistical processes. While a variety of item statistics were reviewed during form building to evaluate the quality of items, one notable statistic that was utilized was differential item functioning (DIF). Items were classified into three categories (A, B, or C) for DIF, ranging from no evidence of DIF to severe DIF, according to the DIF classification convention illustrated in Volume 1. Furthermore, items were categorized positively (i.e., $+\mathrm{A},+\mathrm{B}$, or +C ), signifying that the item favored the focal group (e.g., African-American/Black, Hispanic, or female), or negatively (i.e., $-\mathrm{A},-\mathrm{B}$, or-C), signifying that the item favored the reference group (e.g., White or male). Items were flagged if their DIF statistics indicated the "C" category for any group. A DIF classification of "C" indicates that the item shows significant DIF and should be reviewed for potential content bias, differential validity, or other issues that may reduce item fairness. Items were reviewed by the Bias and Sensitivity Committee regardless of whether the DIF statistic favored the focal or the reference group. The details surrounding this review of items for bias is further described in Volume 2, Test Development.

DIF analyses were conducted for all items to detect potential item bias from a statistical perspective across major ethnic and gender groups. DIF analyses were performed for the following groups:

- Male/Female
- White/African-American
- White/Hispanic
- Not Student with Disability (SWD)/SWD
- Not English language learner (ELL)/ELL

A detailed description of the DIF analysis that was performed is presented in Volume 1, Section 5.2, of the 2017-2018 FSA Annual Technical Report. The DIF statistics for each test item are presented in the appendices of Volume 1 of the 2017-2018 FSA Annual Technical Report.

## Summary

This report is intended to provide a collection of reliability and validity evidence to support appropriate inferences from the observed test scores. The overall results can be summarized as follows:

- Reliability: Various measures of reliability are provided at the aggregate and subgroup levels, showing the reliability of all tests is in line with acceptable industry standards.
- Content validity: Evidence is provided to support the assertion that content coverage on each form was consistent with test specifications of the blueprint across testing modes.
- Internal structural validity: Evidence is provided to support the selection of the measurement model, the tenability of local independence, and the reporting of an overall score and subscores at the reporting category levels.


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[^0]:    * Reporting categories and the number of items belonging to each reporting category are identical for both online and

